

MANE 3332.03

Section 1

MANE 3332.03

1 / 35

MANE 3332.03

Lecture 19, April 3

• Topics:

- Chapter 5: CLT
- Chapter 6: Multivariate Statistical Analysis
- Chapter 7: Definitions
- Chapter 8: Interval Estimation

{ 3 sets of Practice Problems/Quizzes

• Assignments:

- Technical Report One due today
- Linear Combination Practice Problems due today
- Linear Combination Quiz (assigned 4/3/25, due 4/8/25)

• Attendance

• Questions?

2 / 35

Handouts

- Chapter 5
 - Chapter 5 Slides
 - Chapter 5 Slides marked
- Chapter 6
 - Chapter 6 Slides
 - Chapter 6 Slides marked
- Chapter 7
 - Chapter 7 slides
 - Chapter
- Chapter 8
 - Chapter 8 slides
 - Chapter 8 slides marked
- Final Exam Handouts

3 / 35

Class Schedule

Tuesday Lecture	Thursday Lecture
4/1: Chapter 5	4/3: Chapters 7 & 8
4/8: Chapter 8, Case 1	4/10: Chapter 8, Case 2
4/15: Chapter 8, Case 3	4/17: Chapter 9, Case 1
4/22: Chapter 9, Case 2	4/24: Chapter 9, Case 3
4/29: Chapter 11	5/1: Chapter 11
5/6: Review	5/8: Dead Day (no class)

10 Sessions plus final exam

Final Exam: Tuesday May 13, 2025 10:15 am - 12:00 pm

4 / 35

Chapter 8 Introduction

- Intervals are another method of performing estimation
- A confidence interval is an interval estimate on a population parameter (primary focus of this chapter) μ or σ^2
- Three types of interval estimates
 - A confidence intervals bounds population or distribution parameters
 - A tolerance interval bounds a selected proportion of a distribution
 - A prediction interval bounds future observations from the population or distribution
- Interval estimates, especially confidence intervals are commonly used in science and engineering

Tolerance Interval

I am 95% confident that 80% of the distr.
is between 50 and 80.

5 / 35

Chapter 8

MANE 3332.03

Chapter 9

Summary of One-Sample Hypothesis-Testing Procedures

Case	Null Hypothesis	Test Statistic	Alternative Hypothesis	Level Significance Level Criteria for Rejection	P -value	DEC: Curve Parameter	DEC: Curve Shape
1.	$H_0: \mu = \mu_0$ σ^2 known	$z_0 = \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}}$	$H_1: \mu \neq \mu_0$ $H_1: \mu > \mu_0$ $H_1: \mu < \mu_0$	$ z_0 > z_{\alpha/2}$ $z_0 > z_{\alpha}$ $z_0 < -z_{\alpha}$	$P = 2(1 - \Phi(z_0))$ Probability above z_0 $P = 1 - \Phi(z_0)$ $P = \Phi(z_0)$	$d = \mu_0 - \mu_0/\sigma$ $d = (\mu_0 - \mu_0)/\sigma$ $d = (\mu_0 - \mu_0)/\sigma$	e, f e, d e, d
2.	$H_0: \mu = \mu_0$ σ^2 unknown	$t_0 = \frac{\bar{x} - \mu_0}{s/\sqrt{n}}$	$H_1: \mu \neq \mu_0$ $H_1: \mu > \mu_0$ $H_1: \mu < \mu_0$	$ t_0 > t_{\alpha/2, n-1}$ $t_0 > t_{\alpha, n-1}$ $t_0 < -t_{\alpha, n-1}$	Sum of the probability above z_0 and below $-z_0$ Probability above z_0 Probability below z_0	$d = (\mu_0 - \mu_0)/\sigma$ $d = (\mu_0 - \mu_0)/\sigma$ $d = (\mu_0 - \mu_0)/\sigma$	e, f e, d e, d
3.	$H_0: \sigma^2 = \sigma_0^2$ μ unknown	$\chi^2_0 = \frac{(n-1)s^2}{\sigma_0^2}$	$H_1: \sigma^2 \neq \sigma_0^2$ $H_1: \sigma^2 > \sigma_0^2$ $H_1: \sigma^2 < \sigma_0^2$	$\chi^2_0 > \chi^2_{\alpha/2, n-1}$ $\chi^2_0 > \chi^2_{\alpha, n-1}$ $\chi^2_0 < \chi^2_{1-\alpha, n-1}$	See text Section 9.4	$\chi = \sigma_0^2$ $\chi = \sigma_0^2$ $\chi = \sigma_0^2$	e, f e, d e, d
4.	$H_0: p = p_0$ μ, σ^2 unknown	$z_0 = \frac{\hat{p} - p_0}{\sqrt{p_0(1-p_0)}}$	$H_1: p \neq p_0$ $H_1: p > p_0$ $H_1: p < p_0$	$ z_0 > z_{\alpha/2}$ $z_0 > z_{\alpha}$ $z_0 < -z_{\alpha}$	$P = 2(1 - \Phi(z_0))$ Probability above z_0 $P = 1 - \Phi(z_0)$ Probability below z_0	$d = p_0 - p_0$ $d = (p_0 - p_0)$ $d = (p_0 - p_0)$	e, f e, d e, d

Summary of One-Sample Confidence Interval Procedures

Case	Problem Type	Point Estimate	Two-Sided (100- α) Percent Confidence Interval
1.	Mean μ , variance σ^2 known	\bar{x}	$\bar{x} \pm z_{\alpha/2} \sigma/\sqrt{n}$
2.	Mean μ of a normal distribution, variance σ^2 unknown	\bar{x}	$\bar{x} \pm t_{\alpha/2, n-1} s/\sqrt{n}$
3.	Variance σ^2 of a normal distribution	s^2	$\frac{(n-1)s^2}{\chi^2_{\alpha/2, n-1}} < \sigma^2 < \frac{(n-1)s^2}{\chi^2_{1-\alpha/2, n-1}}$
4.	Proportion p parameter of a binomial distribution	\hat{p}	$\hat{p} \pm z_{\alpha/2} \sqrt{\hat{p}(1-\hat{p})}$

6 / 35

Confidence Interval on the Mean of a normal distribution, variance known (Case 1)

- Suppose that X_1, X_2, \dots, X_n is a random sample from a normal population with unknown mean μ and known variance σ^2
- A general expression for a confidence interval is

$$P[L \leq \mu \leq U] = 1 - \alpha$$

- Using the sample results we calculate a $100(1 - \alpha)\%$ confidence of the form

$$l = \bar{x} - z_{\alpha/2} \frac{\sigma}{\sqrt{n}} \leq \mu \leq u = \bar{x} + z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$$

$$l \leq \mu \leq u$$

- A $100(1 - \alpha)\%$ confidence interval for the mean of a normal distribution with variance known is

Impractical case:
usually you
don't know
variance

Case 1 is
statistics are
the easiest

7 / 35

Problem 8-12, part a (6th edition)

8-12. The life in hours of a 75-watt light bulb is known to be normally distributed with $\sigma = 25$ hours. A random sample of 20 bulbs has a mean life of $\bar{x} = 1014$ hours.

(a) Construct a 95% two-sided confidence interval on the mean life.

(b) Construct a 95% lower-confidence bound on the mean life.

Figure 2: image

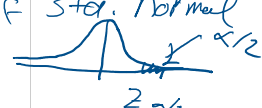
Detail 1) 95% C.I.

$$100(1 - \alpha)\% = 95\%$$

$$(1 - \alpha) = .95$$

$$\alpha = .05$$

Detail 2

$z_{\alpha/2}$
→
upper $\alpha/2$ quantile
of std. normal


Part a) $\bar{x} - z_{\alpha/2} \frac{\sigma}{\sqrt{n}} < \mu < \bar{x} + z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$

$$1014 - 1.96 \frac{25}{\sqrt{20}} \leq \mu < 1014 + 1.96 \frac{25}{\sqrt{20}}$$

$$1014 - 10.75673 \leq \mu \leq 1014 + 10.75673$$

$$1003.24327 \leq \mu \leq 1024.75673$$

8 / 35

Interpreting Confidence Intervals

Montgomery gives the following statement regarding the correct interpretation of confidence intervals.

The correct interpretation lies in the realization that a CI is a random interval because in the probability statement defining the end-points of the interval, L and U are random variables. Consequently, the correct interpretation of a $100(1 - \alpha)\%$ CI depends on the relative frequency view of probability. Specifically, if an infinite number of random samples are collected and a $100(1 - \alpha)\%$ confidence interval for μ is computed from each sample, $100(1 - \alpha)\%$ of these intervals will contain the true value of μ .

9 / 35

One-sided Confidence Bounds

It is possible to construct on-sided confidence bounds

- A $100(1 - \alpha)\%$ upper-confidence bound for μ is

$$\mu \leq u = \bar{x} + z_{\alpha} \frac{\sigma}{\sqrt{n}}$$

- A $100(1 - \alpha)\%$ lower-confidence bound for μ is

$$\bar{x} - z_{\alpha} \frac{\sigma}{\sqrt{n}} = l \leq \mu$$

10 / 35

Sample Size Considerations

If \bar{x} is used as an estimate of μ , we can be $100(1 - \alpha)\%$ confident that the error $|\bar{x} - \mu|$ will not exceed a specified amount E when the sample size is

$$n = \left(\frac{z_{\alpha/2} \sigma}{E} \right)^2$$

11 / 35

Problem 8–12, part b (6th edition)

- 8-12. The life in hours of a 75-watt light bulb is known to be normally distributed with $\sigma = 25$ hours. A random sample of 20 bulbs has a mean life of $\bar{x} = 1014$ hours.
- (a) Construct a 95% two-sided confidence interval on the mean life.
 - (b) Construct a 95% lower-confidence bound on the mean life.

Figure 3: image

12 / 35

A Large Sample CI for μ

- When n is large (say greater than or equal to 40), the central limit theorem can be used
- It states that $\frac{\bar{x} - \mu}{\sigma/\sqrt{n}}$ is approximately a standard normal random variable.
- Thus, we can replace the quantity σ/\sqrt{n} with S/\sqrt{n} and still use the quantiles of the normal distribution to construct a confidence interval.

$$\bar{x} - z_{\alpha/2} \frac{s}{\sqrt{n}} \leq \mu \leq \bar{x} + z_{\alpha/2} \frac{s}{\sqrt{n}}$$

Modification

- What assumption did we relax and why?

for sample data

is from normal distribution

Chapter 8, Case 1 Practice Problems

Confidence Interval for the mean of Normal distribution with variance unknown (Case 2)

The t distribution

- Definition.

Let X_1, X_2, \dots, X_n be a random sample from a normal distribution with unknown mean μ and unknown variance σ^2 . The random variable

$$T = \frac{\bar{X} - \mu}{S/\sqrt{n}}$$

has a t distribution with $n - 1$ degrees of freedom.

- A table of percentage points (quantiles) is given in Appendix A Table 5
- Figure 8-4 on page 180 shows the relationship between the t and normal distributions.
- Figure 8-5 explains the percentage points of the t distribution

15 / 35

Student's t

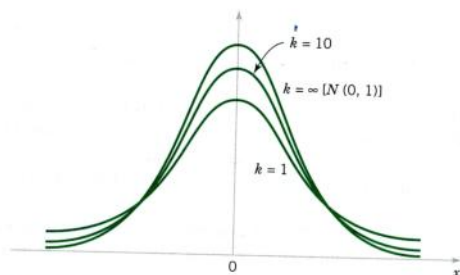


Figure 8-4 Probability density functions of several t distributions.

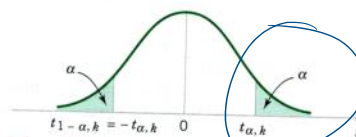


Figure 8-5 Percentage points of the t distribution.

Figure 4: image

16 / 35

Case 1

$$\bar{X} - Z_{\alpha/2} \frac{\sigma}{\sqrt{n}} \leq \mu \leq \bar{X} + Z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$$

Confidence interval definition

- Using the t distribution it is possible to construct CIs

If \bar{x} and s are the mean and standard deviation of a random sample from a normal distribution with unknown variance σ^2 , a $100(1 - \alpha)\%$ confidence interval on μ is given by

$$\hat{\sigma} = s$$

$$\bar{x} - t_{\alpha/2, n-1} \frac{s}{\sqrt{n}} \leq \mu \leq \bar{x} + t_{\alpha/2, n-1} \frac{s}{\sqrt{n}}$$

where $t_{\alpha/2, n-1}$ is the upper $100(\alpha/2)$ percentage point of the t distribution with $n - 1$ degrees of freedom.

17 / 35

Problem 8-30 (6th edition)

8-30. An article in *Nuclear Engineering International* (February 1988, p. 33) describes several characteristics of fuel rods used in a reactor owned by an electric utility in Norway. Measurements on the percentage of enrichment of 12 rods were reported as follows:

2.94	3.00	2.90	2.75	3.00	2.95
2.90	2.75	2.95	2.82	2.81	3.05

- Use a normal probability plot to check the normality assumption.
- Find a 99% two-sided confidence interval on the mean percentage of enrichment. Are you comfortable with the statement that the mean percentage of enrichment is 2.95 percent? Why?

Figure 5: image

99% C.I.

$$1 - \alpha = .99$$

$$\alpha = .01$$

$$n = 12$$

$$\bar{x} = 2.902$$

$$s = .0993$$

$$\bar{x} - t_{\alpha/2, n-1} \frac{s}{\sqrt{n}} \leq \mu \leq \bar{x} + t_{\alpha/2, n-1} \frac{s}{\sqrt{n}}$$

$$2.902 - t_{.005, 11} \frac{.0993}{\sqrt{12}}$$

$$\leq \mu$$

$$\leq$$

$$2.902 + t_{.005, 11} \frac{.0993}{\sqrt{12}}$$

$$t_{.005, 11} = 3.106 \quad 2.902 - 3.106 \frac{.0993}{\sqrt{12}}$$

$$\leq \mu$$

$$\leq 2.902 + 3.106 \frac{.0993}{\sqrt{12}}$$

$$2.902 \leq \mu \leq 2.991$$

18 / 35

Table IV Percentage Points $t_{\alpha, v}$ of the t -Distribution

α	.40	.25	.10	.05	.025	.01	.005	.0025	.001	.0005
1	.325	1.000	3.078	6.314	12.706	31.821	63.657	127.32	318.31	636.62
2	.289	.816	1.886	2.920	4.303	6.965	9.925	14.089	23.326	31.598
3	.277	.765	1.638	2.353	3.182	4.541	5.841	7.453	10.213	12.924
4	.271	.741	1.533	2.132	2.776	3.747	4.604	5.598	7.173	8.610
5	.267	.727	1.476	2.015	2.571	3.365	4.032	4.773	5.893	6.869
6	.265	.718	1.440	1.943	2.447	3.143	3.707	4.317	5.208	5.959
7	.263	.711	1.415	1.895	2.365	2.998	3.499	4.029	4.785	5.408
8	.262	.706	1.397	1.860	2.306	2.896	3.355	3.833	4.501	5.041
9	.261	.703	1.383	1.833	2.262	2.821	3.250	3.690	4.297	4.781
10	.260	.700	1.372	1.812	2.228	2.764	3.169	3.581	4.144	4.587
11	.260	.697	1.363	1.796	2.201	2.718	3.106	3.497	4.025	4.437
12	.259	.695	1.356	1.782	2.179	2.681	3.055	3.428	3.930	4.318
13	.259	.694	1.350	1.771	2.160	2.650	3.012	3.372	3.852	4.221
14	.258	.692	1.345	1.761	2.145	2.624	2.977	3.326	3.787	4.140
15	.258	.691	1.341	1.753	2.131	2.602	2.947	3.286	3.733	4.073
16	.258	.690	1.337	1.746	2.120	2.583	2.921	3.252	3.686	4.015
17	.257	.689	1.333	1.740	2.110	2.567	2.898	3.222	3.646	3.963
18	.257	.688	1.330	1.734	2.101	2.552	2.878	3.197	3.610	3.921
19	.257	.688	1.328	1.729	2.093	2.539	2.861	3.174	3.579	3.883
20	.257	.687	1.325	1.725	2.086	2.528	2.845	3.153	3.552	3.850
21	.257	.686	1.323	1.721	2.080	2.518	2.831	3.135	3.527	3.819
22	.256	.686	1.321	1.717	2.074	2.508	2.819	3.119	3.505	3.792
23	.256	.685	1.319	1.714	2.069	2.500	2.807	3.104	3.485	3.767
24	.256	.685	1.318	1.711	2.064	2.492	2.797	3.091	3.467	3.745
25	.256	.684	1.316	1.708	2.060	2.485	2.787	3.078	3.450	3.725
26	.256	.684	1.315	1.706	2.056	2.479	2.779	3.067	3.435	3.707
27	.256	.684	1.314	1.703	2.052	2.473	2.771	3.057	3.421	3.690
28	.256	.683	1.313	1.701	2.048	2.467	2.763	3.047	3.408	3.674
29	.256	.683	1.311	1.699	2.045	2.462	2.756	3.038	3.396	3.659
30	.256	.683	1.310	1.697	2.042	2.457	2.750	3.030	3.385	3.646
40	.255	.681	1.303	1.684	2.021	2.423	2.704	2.971	3.307	3.551
60	.254	.679	1.296	1.671	2.000	2.390	2.660	2.915	3.232	3.460
120	.254	.677	1.289	1.658	1.980	2.358	2.617	2.860	3.160	3.373
∞	.253	.674	1.282	1.645	1.960	2.326	2.576	2.807	3.090	3.291

 v = degrees of freedom.

Upper Bound
 ~~$X \leq \mu \pm k$~~

$$\mu \leq \bar{X} + t_{\alpha, n-1} \frac{s}{\sqrt{n}}$$

Lower Bound
 ~~$\mu \leq \mu \pm k$~~

$$\bar{X} - t_{\alpha, n-1} \frac{s}{\sqrt{n}} \leq \mu$$

One-sided confidence bounds

- Are easy to construct
- Use only the appropriate upper or lower bound
- Change $t_{\alpha/2, n-1}$ to $t_{\alpha, n-1}$

Chapter 8, Case 2 Practice Problems

21 / 35

Confidence Interval for σ^2 and σ (Case 3)

- Section 8-3 presents a CI for σ^2 or σ
- Requires the χ^2 (chi-squared) distribution

Let X_1, X_2, \dots, X_n be a random sample from a normal distribution with mean μ and variance σ^2 and let S^2 be the sample variance. Then the random variable

$$\chi^2 = \frac{(n-1)S^2}{\sigma^2}$$

has a chi-square (χ^2) distribution with $n - 1$ degrees of freedom

- A table of the upper percentage points of the χ^2 distribution are given in Table 4 in the appendix
- Figure 8-9 on page 183 explains the percentage points of the χ^2 distribution

22 / 35

$$S^2 = \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n-1}$$

χ^2 - chi

Always work
 σ^2 and if
 σ is desired
take $\sqrt{\quad}$

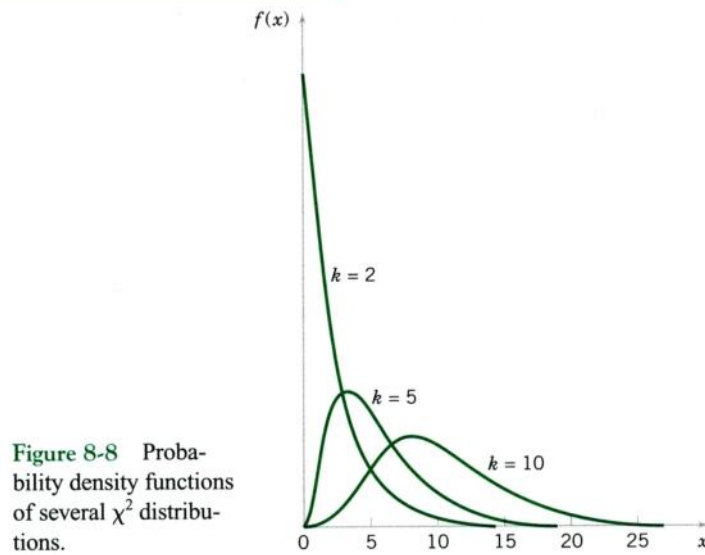


Figure 7: image

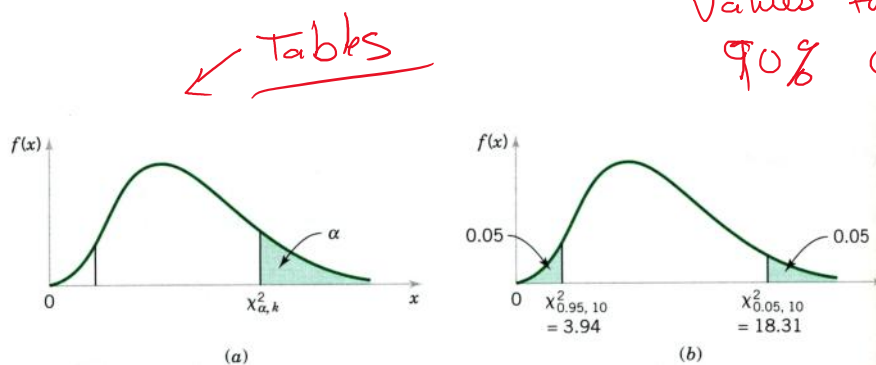
Figure 8-9 Percentage point of the χ^2 distribution. (a) The percentage point $\chi^2_{\alpha,k}$. (b) The upper percentage point $\chi^2_{0.05,10} = 18.31$ and the lower percentage point $\chi^2_{0.95,10} = 3.94$.

Figure 8: image

Confidence Intervals for σ^2 and σ

If s^2 is the sample variance from a random sample of n observations from a normal distribution with unknown variance σ^2 , then a $100(1 - \alpha)\%$ confidence interval on σ^2 is

$$l = \frac{(n-1)s^2}{\chi_{\alpha/2, n-1}^2} \leq \sigma^2 \leq \frac{(n-1)s^2}{\chi_{1-\alpha/2, n-1}^2} = u$$

where $\chi_{\alpha/2, n-1}^2$ and $\chi_{1-\alpha/2, n-1}^2$ are the upper and lower $100\alpha/2$ percentage points of the χ^2 -distribution with $n - 1$ degrees of freedom

25 / 35

Problem 8-36 (6th edition)

8-36. The sugar content of the syrup in canned peaches is normally distributed. A random sample of $n = 10$ cans yields a sample standard deviation of $s = 4.8$ milligrams. Find a 95% two-sided confidence interval for σ .

Figure 9: image

Given $n = 10$
 $S = 4.8$
 $\alpha = .05$

Need $\chi_{.025, 9}^2 = 19.02$

$\chi_{.975, 9}^2 = 2.70$

$$\frac{(n-1)S^2}{\chi_{\alpha/2, n-1}^2} \leq \sigma^2 \leq \frac{(n-1)S^2}{\chi_{1-\alpha/2, n-1}^2}$$

$$\frac{(10-1)(4.8)^2}{19.02} \leq \sigma^2 \leq \frac{(10-1)(4.8)^2}{2.70}$$

$$10.902 \leq \sigma^2 \leq 76.8$$

$$\sqrt{10.902} \leq \sigma \leq \sqrt{76.8}$$

$$3.3018 \leq \sigma \leq 8.764$$

Table III Percentage Points χ^2_{α} of the Chi-Squared Distribution

α	.995	.990	.975	.950	.900	.800	.700	.600	.500	.400	.300	.200	.100	.050	.025	.010	.005
1	.004	.006	.010	.016	.020	.025	.030	.035	.040	.045	.050	.055	.060	.065	.070	.075	.080
2	.010	.014	.020	.026	.032	.038	.044	.050	.056	.062	.068	.074	.080	.086	.092	.098	.104
3	.016	.020	.026	.032	.038	.044	.050	.056	.062	.068	.074	.080	.086	.092	.098	.104	.110
4	.020	.024	.030	.036	.042	.048	.054	.060	.066	.072	.078	.084	.090	.096	.102	.108	.114
5	.024	.028	.034	.040	.046	.052	.058	.064	.070	.076	.082	.088	.094	.100	.106	.112	.118
6	.028	.032	.038	.044	.050	.056	.062	.068	.074	.080	.086	.092	.098	.104	.110	.116	.122
7	.032	.036	.042	.048	.054	.060	.066	.072	.078	.084	.090	.096	.102	.108	.114	.120	.126
8	.036	.040	.046	.052	.058	.064	.070	.076	.082	.088	.094	.100	.106	.112	.118	.124	.130
9	.040	.044	.050	.056	.062	.068	.074	.080	.086	.092	.098	.104	.110	.116	.122	.128	.134
10	.044	.048	.054	.060	.066	.072	.078	.084	.090	.096	.102	.108	.114	.120	.126	.132	.138
11	.048	.052	.058	.064	.070	.076	.082	.088	.094	.100	.106	.112	.118	.124	.130	.136	.142
12	.052	.056	.062	.068	.074	.080	.086	.092	.098	.104	.110	.116	.122	.128	.134	.140	.146
13	.056	.060	.066	.072	.078	.084	.090	.096	.102	.108	.114	.120	.126	.132	.138	.144	.150
14	.060	.064	.070	.076	.082	.088	.094	.100	.106	.112	.118	.124	.130	.136	.142	.148	.154
15	.064	.068	.074	.080	.086	.092	.098	.104	.110	.116	.122	.128	.134	.140	.146	.152	.158
16	.068	.072	.078	.084	.090	.096	.102	.108	.114	.120	.126	.132	.138	.144	.150	.156	.162
17	.072	.076	.082	.088	.094	.100	.106	.112	.118	.124	.130	.136	.142	.148	.154	.160	.166
18	.076	.080	.086	.092	.098	.104	.110	.116	.122	.128	.134	.140	.146	.152	.158	.164	.170
19	.080	.084	.090	.096	.102	.108	.114	.120	.126	.132	.138	.144	.150	.156	.162	.168	.174
20	.084	.088	.094	.100	.106	.112	.118	.124	.130	.136	.142	.148	.154	.160	.166	.172	.178
21	.088	.092	.098	.104	.110	.116	.122	.128	.134	.140	.146	.152	.158	.164	.170	.176	.182
22	.092	.096	.102	.108	.114	.120	.126	.132	.138	.144	.150	.156	.162	.168	.174	.180	.186
23	.096	.100	.106	.112	.118	.124	.130	.136	.142	.148	.154	.160	.166	.172	.178	.184	.190
24	.100	.104	.110	.116	.122	.128	.134	.140	.146	.152	.158	.164	.170	.176	.182	.188	.194
25	.104	.108	.114	.120	.126	.132	.138	.144	.150	.156	.162	.168	.174	.180	.186	.192	.198
26	.108	.112	.118	.124	.130	.136	.142	.148	.154	.160	.166	.172	.178	.184	.190	.196	.202
27	.112	.116	.122	.128	.134	.140	.146	.152	.158	.164	.170	.176	.182	.188	.194	.200	.206
28	.116	.120	.126	.132	.138	.144	.150	.156	.162	.168	.174	.180	.186	.192	.198	.204	.210
29	.120	.124	.130	.136	.142	.148	.154	.160	.166	.172	.178	.184	.190	.196	.202	.208	.214
30	.124	.128	.134	.140	.146	.152	.158	.164	.170	.176	.182	.188	.194	.200	.206	.212	.218
40	.149	.153	.159	.165	.171	.177	.183	.189	.195	.201	.207	.213	.219	.225	.231	.237	.243
50	.175	.179	.185	.191	.197	.203	.209	.215	.221	.227	.233	.239	.245	.251	.257	.263	.269
60	.200	.204	.210	.216	.222	.228	.234	.240	.246	.252	.258	.264	.270	.276	.282	.288	.294
70	.224	.228	.234	.240	.246	.252	.258	.264	.270	.276	.282	.288	.294	.300	.306	.312	.318
80	.248	.252	.258	.264	.270	.276	.282	.288	.294	.300	.306	.312	.318	.324	.330	.336	.342
90	.272	.276	.282	.288	.294	.300	.306	.312	.318	.324	.330	.336	.342	.348	.354	.360	.366
100	.296	.300	.306	.312	.318	.324	.330	.336	.342	.348	.354	.360	.366	.372	.378	.384	.390

 ν = degrees of freedom.

One-sided confidence bounds

- Are easy to construct
- Use only the appropriate upper or lower bound
- Change $\chi^2_{\alpha/2, n-1}$ to $\chi^2_{\alpha, n-1}$ or $\chi^2_{1-\alpha/2, n-1}$ to $\chi^2_{1-\alpha, n-1}$
- See eqn (8-20) on page 184

Lower Conf. Bound

$$\frac{(n-1)s^2}{\chi^2_{\alpha, n-1}} \leq \sigma^2$$

Upper Bound

$$\sigma^2 \leq \frac{(n-1)s^2}{\chi^2_{1-\alpha, n-1}}$$

Chapter 8, Case 3 Practice Problems

29 / 35

Binomial(p)

Large-Sample CI for a Population Proportion (Case 4)

Not Covered

- Recall from chapter 4, that the sampling distribution of \hat{P} is approximately normal with mean p and variance $p(1-p)/n$, if n is not too close to either 0 or 1 and if n is relatively large.
- Typically, we require both $np \geq 5$ and $n(1-p) \geq 5$

If n is large, the distribution of

$$Z = \frac{\hat{P} - p}{\sqrt{\frac{p(1-p)}{n}}} = \frac{\hat{P} - p}{\sqrt{np(1-p)}/n}$$

is approximately standard normal.

30 / 35

If \hat{p} is the proportion of observations in a random sample of size n that belongs to a class of interest, an approximate $100(1 - \alpha)\%$ confidence interval on the proportion p of the population that belongs to this class is

$$\hat{p} - z_{\alpha/2} \sqrt{\frac{\hat{p}(1 - \hat{p})}{n}} \leq p \leq \hat{p} + z_{\alpha/2} \sqrt{\frac{\hat{p}(1 - \hat{p})}{n}}$$

where $z_{\alpha/2}$ is the upper $\alpha/2$ percentage point of the standard normal distribution

Other Considerations

- We can select a sample so that we are $100(1 - \alpha)\%$ confident that error $E = |p - \hat{P}|$ using

$$n = \left(\frac{z_{\alpha/2}}{E} \right)^2 p(1 - p)$$

- An upper bound on is given by

$$n = \left(\frac{z_{\alpha/2}}{E} \right)^2 (0.25)$$

- One-sided confidence bounds are given in eqn (8-26) on page 187

Guidelines for Constructing Confidence Intervals

- Review excellent guide given in Table 8-1

Other Interval Estimates

- When we want to predict the value of a single value in the future, a **prediction interval** is used
- A **tolerance interval** captures $100(1 - \alpha)\%$ of observations from a distribution

33 / 35

Not Covered

Prediction Interval for a Normal Distribution

- Excellent discussion on pages 189 - 190
- A $100(1 - \alpha)\%$ PI on a single future observation from a normal distribution is given by
-

$$\bar{x} - t_{\alpha/2, n-1} s \sqrt{1 + \frac{1}{n}} \leq X_{n+1} \leq \bar{x} + t_{\alpha/2, n-1} s \sqrt{1 + \frac{1}{n}}$$

34 / 35

Not Covered

Tolerance Intervals for a Normal Distribution

- A **tolerance interval** to contain at least $\gamma\%$ of the values in a normal population with confidence level $100(1 - \alpha)\%$ is

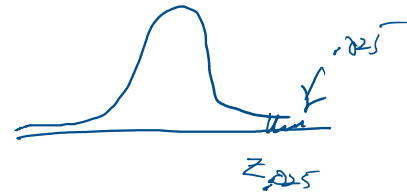
$$\bar{x} - ks, \bar{x} + ks$$

where k is a tolerance interval factor for the normal distribution found in appendix A Table XII. Values are given for $1 - \alpha = 0.9, 0.95$ and 0.99 confidence levels and for $\gamma = .90, .95$, and $.99$ probability of coverage

- One-sided tolerance bounds can also be computed. The factors are also in Table XII

95% C.I. $\rightarrow \alpha = .05$

$$Z_{\alpha/2} = Z_{.05/2} = Z_{.025}$$



Easiest way to find $Z_{\alpha/2}$ is to use the t-tables

$$Z_{.025} = t_{\infty, .025} = 1.96$$

Confidence bound

Tuesday, April 8, 2025 11:28 AM

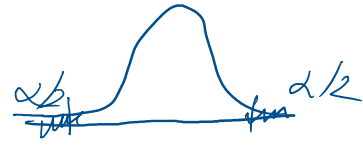
$$100(1-\alpha)\% \text{ C.I.}$$

$$l \leq \mu \leq u$$

$$\bar{X} - z_{\alpha/2} \sigma/\sqrt{n} \leq \mu \leq \bar{X} + z_{\alpha/2} \sigma/\sqrt{n}$$

Lower Bound

$$\bar{X} - z_{\alpha} \sigma/\sqrt{n} \leq \mu \leq \bar{X}$$



Upper Bound

$$\bar{X} \leq \mu \leq \bar{X} + z_{\alpha} \sigma/\sqrt{n}$$

95% Lower Confidence on μ

$$\bar{y} = 1014, \sigma = 25, n = 20$$

$$\bar{X} - Z_{\alpha} \sigma / \sqrt{n} \leq \mu$$

$$1014 - Z_{.05} \frac{25}{\sqrt{20}} \leq \mu$$

$$\text{Recall } t_{\infty, .05} = Z_{.05}$$

$$1014 - 1.645 \frac{25}{\sqrt{20}} \leq \mu$$

$$\underline{1004.80417 \leq \mu}$$

$$\mu \leq \bar{X}$$

$$100(1-\alpha)\% = 95\%$$

$$(1-\alpha) = .95$$

$$\alpha = .05$$

Chapter 8, Case 1 PP

Tuesday, April 8, 2025 11:46 AM

QUESTION 1

Consider a sample of size 15 from a normal distribution with mean \bar{X} and sigma σ 15.69. What is the value of a two-sided 99.9% confidence interval for the mean?

Screen clipping taken: 4/8/2025 11:46 AM

$$100(1-\alpha)\% = 99.9$$

$$1-\alpha = .999$$

$$\alpha = .001$$

$$\bar{X} - Z_{\alpha/2} \sigma / \sqrt{n} \leq \mu \leq \bar{X} + Z_{\alpha/2} \sigma / \sqrt{n}$$

$$Z_{.001/2} = Z_{.0005} = t_{\infty, .0005} = 3.291$$

Chapter 8, Case 1 PP

Tuesday, April 8, 2025 11:49 AM

QUESTION 4

Consider a sample of size 10 from a normal distribution with mean 16.7 and sigma 3.65. What is the value of a 90.0 % upper-confidence bound for the mean?

- ☐ mu <= 18.599
- ☐ mu <= 18.407
- ☐ mu <= 18.891
- ☐ mu <= 17.168
- ☐ mu <= 22.099
- ☐ mu <= 18.179

Screen clipping taken: 4/8/2025 11:49 AM

$$100(1-\alpha)\% = 90\%$$

$$(1-\alpha) = .9$$

$$\alpha = .1$$

$$\cancel{L} \leq \mu \leq U$$

$$U \leq \bar{X} + Z_{\alpha} \sigma / \sqrt{n}$$

$$\text{need } Z_{.1} = t_{\infty, .1} = 1.282$$

Chapter 8, Case 1 PP

Tuesday, April 8, 2025 11:52 AM

QUESTION 5

Consider a sample of size 30 from a normal distribution with mean 30.8 and sigma 3.85. What is the value of a 99.0 % lower-confidence bound for the mean?

- ☐ mu >= 24.504
- ☐ mu >= 29.651
- ☐ mu >= 29.527
- ☐ mu >= 30.501
- ☐ mu >= 28.989
- ☐ mu >= 29.165

Screen clipping taken: 4/8/2025 11:52 AM

$$\bar{X} - Z_{\alpha} \sigma / \sqrt{n} \leq \mu$$

$$Z_{.01} = t_{\infty, .01} = 2.326$$

$$100(1-\alpha)\% = 99\%$$

$$(1-\alpha) = .99$$

$$\alpha = .01$$

40 on midterm

$$-60\left(\frac{1}{4}\right) = -15$$

97.5% C.B. $\rightarrow \alpha$

QUESTION 1

Consider a sample of size 27 from a normal distribution with mean 12.3 and s 2.45. What is the value of a 97.5 % lower-confidence bound for the mean?

- ☐ mu >= 9.925
- ☐ mu >= 11.843
- ☒ mu >= 11.331
- ☐ mu >= 11.771
- ☐ mu >= 11.178
- ☐ mu >= 12.113

$$100(1-\alpha)\% = 97.5\%$$

$$(1-\alpha) = .975$$

$$\alpha = \underline{.025}$$

\bar{x}

Sample Standard deviation

$$\bar{X} - t_{\alpha, n-1} \frac{s}{\sqrt{n}} \leq \mu$$

$$\text{need } t_{.025, 27-1} = 2.056$$

$$12.3 - 2.056 \frac{2.45}{\sqrt{27}} \leq \mu$$

$$11.33059 \leq \mu$$

Screen clipping taken: 4/10/2025 11:29 AM

Chapter 8, case 2 pp

Thursday, April 10, 2025 11:33 AM

$$\text{So } \alpha = 1 - 0.9975 = 0.0025$$

QUESTION 2

Consider a sample of size 7 from a normal distribution with mean 17.0 and s 3.32. What is the value of a 99.75 % upper-confidence bound for the mean?

☐ $\mu \leq 19.047$

☐ $\mu \leq 24.843$

Screen clipping taken: 4/10/2025 11:33 AM

$$\mu \leq \bar{X} + t_{\alpha, n-1} \frac{s}{\sqrt{n}}$$

$$\text{need } t_{0.0025, 7-1} = 4.317$$

$$\mu \leq 17 + 4.317 \left(\frac{3.32}{\sqrt{7}} \right)$$

$$\mu \leq 22.41715$$

QUESTION 5

Consider a sample of size 23 from a normal distribution with mean 25.7 and $s = 6.19$. What is the value of a two-sided 99.8 % confidence interval for the mean?

- ☐ (20.346, 31.054)
- ☐ (21.552, 29.848)
- ☐ (21.176, 30.224)
- ☐ (24.757, 26.643)
- ☐ (-2.303, 53.703)
- ☐ (19.861, 31.539)

Screen clipping taken: 4/10/2025 11:37 AM

1) Find α $\rightarrow 1 - \alpha = .998$, so $\alpha = .002$

2) need $t_{\alpha/2, n-1} = t_{.001, 22} = 3.505$

$$\bar{X} - t_{\alpha/2, n-1} \frac{s}{\sqrt{n}} \leq \mu \leq \bar{X} + t_{\alpha/2, n-1} \frac{s}{\sqrt{n}}$$

$$25.7 - 3.505 \left(\frac{6.19}{\sqrt{23}} \right) \leq \mu \leq 25.7 + 3.505 \left(\frac{6.19}{\sqrt{23}} \right)$$

$$21.17608 \leq \mu \leq 30.22392$$

Attendance 1-D

Chapter 8, Case 3

Tuesday, April 15, 2025 11:20 AM

QUESTION 1

Consider a sample of size n from a normal distribution with mean \bar{x} and sample standard deviation s . What is the value of lower 90.0 % confidence bound on the variance?

☐ 58.381

☒ 66.489

☐ 57.778

☐ 61.937

☐ 8.154

Screen clipping taken: 4/15/2025 11:20 AM

$$(1 \leq \sigma^2) \text{ is } \text{not correct}$$

$$\text{need } \alpha = .10$$

$$\frac{(n-1)s^2}{\chi^2_{\alpha, n-1}} \leq \sigma^2$$

$$\chi^2_{.1, 12-1} = 17.28$$

$$\frac{(12-1)(10.22)^2}{17.28} \leq \sigma^2$$

$$66.489 \leq \sigma^2$$

Chapter 8, case 3

Tuesday, April 15, 2025 11:26 AM

QUESTION 2

Consider a sample of size 24 from a normal distribution with mean 61.4 and sample standard deviation 16.71. What is the value of a two-sided 90.0 % confidence interval for the variance?

- ☐ (55.554, 67.246)
- ☐ (200.63, 432.468)
- ☒ (182.603, 490.615)
- ☐ (13.513, 22.15)
- ☐ (176.336, 463.693)

Screen clipping taken: 4/15/2025 11:26 AM

$$\alpha = .1$$

need

$$\chi^2_{\alpha/2, n-1} = \chi^2_{.05, 23} = 35.17$$

$$\chi^2_{1-\alpha/2, n-1} = \chi^2_{.95, 23} = 13.09$$

$$s = 16.71$$

n

$$\frac{(n-1)s^2}{\chi^2_{\alpha/2, n-1}} \leq \sigma^2 \leq \frac{(n-1)s^2}{\chi^2_{1-\alpha/2, n-1}}$$

$$\frac{(24-1)(16.71)^2}{\chi^2_{.05, 23}} \leq \sigma^2 \leq \frac{(24-1)(16.71)^2}{\chi^2_{.95, 23}}$$

$$\frac{(24-1)(16.71)^2}{35.17} \leq \sigma^2 \leq \frac{(24-1)(16.71)^2}{13.09}$$

$$182.603 \leq \sigma^2 \leq 490.615$$

Chapter 8, Case 3

Tuesday, April 15, 2025 11:34 AM

QUESTION 5

Consider a sample of size 8 from a normal distribution with mean 27.1 and sample standard deviation 10.19. What is the value of an upper 90.0 % confidence bound on the standard deviation?

- ☐ 32.198
- ☐ 18.302
- ☐ 256.838
- ☒ 16.026
- ☐ 14.431

Screen clipping taken: 4/15/2025 11:34 AM

$$s = 10.19$$

$$n = 8$$

$$\sigma$$

$$\sigma^2 \leq \frac{(n-1)s^2}{\chi^2_{1-\alpha, n-1}}$$

$$\sigma^2 \leq \frac{(8-1)(10.19)^2}{2.83} = 256.838$$

$$\text{So } \sigma < \sqrt{256.838} = 16.026$$

Know

$$n = 8$$

$$s = 10.19$$

$$\alpha = .1$$

Need $\chi^2_{1-.1, 8-1}$

$$= \chi^2_{.9, 7}$$

$$= 2.83$$