

Probability Density Function

- Notice the difference from a discrete random variable
- The formal definition of a probability density function is a function such that such that

 - of $f(x) \ge 0$ of $\int_{-\infty}^{\infty} f(x) dx = 1$ of $P(a \le X \le b) = \int_{a}^{b} f(x) dx$

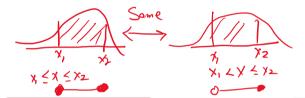
Probability Density Function

Any interesting property of continuous random variables is

$$P(x_1 \le X \le x_2) = P(x_1 < X \le x_2) = P(x_1 \le X < x_2)$$

= $P(x_1 < X < x_2)$

- Does not apply to discrete random variables
- Explanation





Cumulative Distribution Function

The cumulative distribution function for a continuous random variable \boldsymbol{X} is

$$F(x) = P(X \le x) = \int_{-\infty}^{x} f(y) \, dy$$

6 / 42

Mean and Variance of a Continuous Random Variable



• The mean value of a continuous Random Variable chopton

$$\mu = E(X) = \int_{-\infty}^{\infty} x f(x) dx \qquad \text{if } \sum_{\text{ally}} x f(y)$$

• The variance of a continuous random variable is defined to be

$$\sigma^{2} = V(X) = E(X - \mu)^{2}$$
$$= \int_{-\infty}^{\infty} (x - \mu)^{2} f(x) dx = \int_{-\infty}^{\infty} x^{2} f(x) dx - \mu^{2}$$

ullet The standard deviation of X is

$$\sigma = \sqrt{V(X)}$$

Continuous Uniform Distribution

The continuous uniform distribution is the analog of the discrete uniform distribution in that all outcomes are equally likely to occur

ullet A continuous uniform distribution for the random variable X has a probability density function

$$f(x) = \frac{1}{b-a}, \ a \le x \le b$$

$$f(x) = \frac{1}{b-a}, \quad a \le x \le b$$
• The mean of the uniform distribution is
$$\mu = E(X) = \frac{a+b}{2}$$
• The variance of X is
$$\mu = E(X) = \frac{a+b}{2}$$
• The variance of X is

$$\sigma^{2} - V(x) - \frac{(b-a)^{2}}{b-a} = \frac{1}{b-a} \int x dx$$

$$= \left(\frac{1}{b-a}\right) \frac{x^{2}}{2} \begin{vmatrix} b \\ x-a \end{vmatrix}$$

$$= \left(\frac{1}{b-a}\right) \left(\frac{x^{2}}{2} - \frac{a^{2}}{2}\right)$$

$$= \frac{1}{2} \left(\frac{1}{b-a}\right) (a+b)(a-b-a)$$

f(x)= 2 for 4975 x x x 50.25 $P(x) = \int_{50}^{5005} 2 dx = 2x \Big|_{x=50}^{5005}$ Uniform Problem 4.1.6 =2(50.25-50)=.5 $P(X \angle 49.8) = \int_{49.75}^{49.8} 2dx = 2x |_{y=49.75}$ = 2 (49.8 - 49.75) = .1• See page P-25 P(x249.8) = F(49.8) P(x>50) = 1-F(50) Affendance 1-D



The Normal Distribution

- The normal distribution is the most widely used and important distribution in statistics.
- You must master this!
- A random variable X with probability density function

$$f(x) = \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$
 for $-\infty < x < \infty$

has a normal distribution with parameters μ and σ where $-\infty < \mu < \infty$

- and $\sigma>0$ The normal distribution with parameters μ and σ is denoted $N(\mu,\sigma^2)$
- An interesting web-site is http://www.seeingstatistics.com/seeingTour/wormal/shape3.html

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Symmetric

J2 < 9

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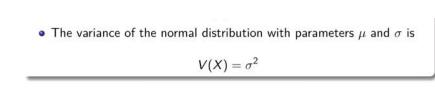
Mean and Variance of the Normal Distribution

ullet The mean of the normal distribution with parameters μ and σ is

$$E(X) = \mu$$

 \bullet The variance of the normal distribution with parameters μ and σ is

$$V(X) - \sigma^2$$



11/42

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Central Limit Theorem

- Brief introduction
- States that the distribution of the average of independent random variables will tend towards a normal distribution as n gets large
- More details later

mgets large

Calculating Normal Probabilities

- Is somewhat complicated
- The difficulty is $\int_a^b f(x) dx$ does not have a closed form solution
- Probabilities must be found by numerical techniques (tabled values)
- It is very helpful to draw a sketch of the desired probabilities (I require this)

How to use without closed form Solution () Tables (2) Compater

Closed form reall uniform

If a) du = X-9

a ba

answer does not contain

an integral

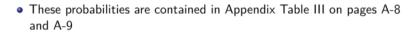
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Chanter A Content

The Standard Normal Distribution

- A normal random variable with $\mu=0$ and $\sigma=1$ is called a **standard** normal random variable
- A standard normal random variable is denoted as z
- The cumulative distribution function for a standard normal is defined to be the function (x) for standard normal is defined to be the function (x)

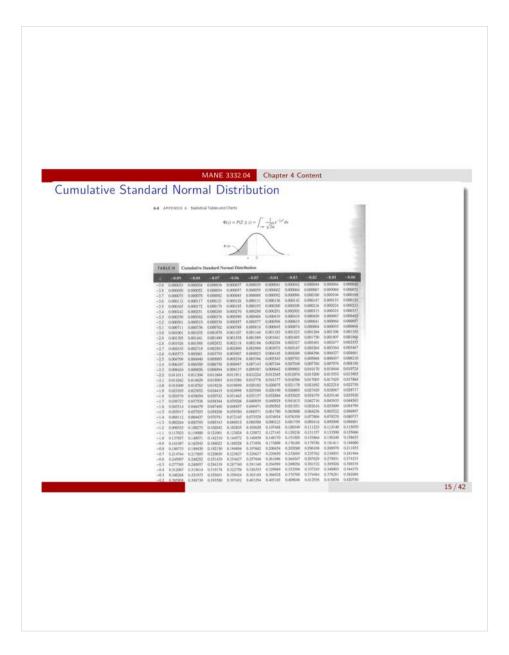
$$\Phi(z) = P(Z \leq z) = F(z)$$

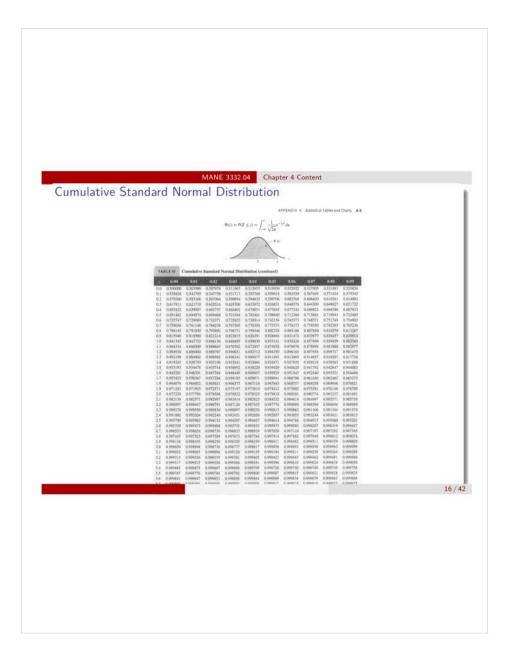


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Chapter 4 Page 13





Standard Normal Problem

5.1.1 Suppose that $Z \sim N(0, 1)$. Find:

- (a) $P(Z \le 1.34)$
- (b) $P(Z \ge -0.22)$
- (c) $P(-2.19 \le Z \le 0.43)$
- (d) $P(0.09 \le Z \le 1.76)$
- (e) $P(|Z| \le 0.38)$
- (f) The value of x for which $P(Z \le x) = 0.55$
- (g) The value of x for which $P(Z \ge x) = 0.72$
- (h) The value of x for which $P(|Z| \le x) = 0.31$

Figure 3: image



Normal pdf
$$f(x) = \frac{1}{12\pi^{2}\sigma} \exp\left[-\frac{1}{2}\left(\frac{x-\nu}{\sigma}\right)^{2}\right]$$

$$= \frac{1}{12\pi^{2}\sigma} e^{-\frac{1}{2}\left(\frac{x-\nu}{\sigma}\right)^{2}}$$

Standardizing (the z-transform)

ullet Suppose X is a normal random variable with mean μ and variance σ^2

Y-N Trons forms to Standard normal $P(X \le x) = P\left(\frac{X - \mu}{\sigma} \le \frac{x - \mu}{\sigma}\right) = P(Z \le z)$

- The z-value is $z = (x \mu)/\sigma$
- Result allows the standard normal tables to be used to calculate probabilities for any normal distribution

Data $P(x \le 10.34)$ $P(x \le 10.34) = P(2 \le \frac{10.34 - 10}{2})$ $= P(2 \le .17) = 0.17$ = .567 495MANE 3332.04 Chapter 4 Content Normal Probability Problem N=10,0=2 5.1.3 Suppose that $X \sim N(10, 2)$. Find: (a) $P(X \le 10.34)$ (b) $P(X \ge 11.98)$ (c) $P(7.67 \le X \le 9.90)$ (d) $P(10.88 \le X \le 13.22)$ (e) $P(|X - 10| \le 3)$ (f) The value of x for which $P(X \le x) = 0.81$ (g) The value of x for which $P(X \ge x) = 0.04$ (h) The value of x for which $P(|X - 10| \ge x) = 0.63$

Figure 4: image

Normal Practice Problems

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Normal Approximation to the Binomial Distribution

• If X is a binomial random variable,

$$Z = \frac{X - np}{\sqrt{np(1 - p)}}$$

is approximately a standard normal random variable. Consequently, probabilities computed from Z can be used to approximate probabilities for

- II----II-- I--II----I----

naive 7= X-1

is approximately a standard normal random variable. Consequently, probabilities computed from Z can be used to approximate probabilities for Χ

Usually holds when

$$np > 5$$
 and $n(1-p) > 5$

22 / 42

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Problem

- 4. A process yields 10% defective items. If 100 items are randomly selected from the process, what is the probability that the number of defectives
- (a) exceeds 13?
- (b) is less than 8?

Source: Walpole, Myers, Myers & Ye

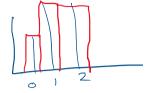
Figure 5: image

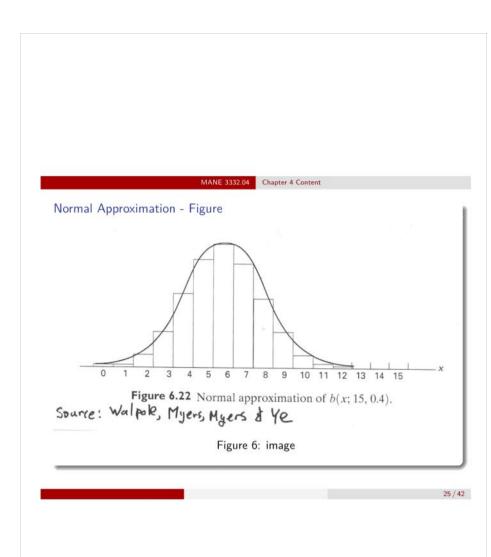
• How good are the approximations?

Continuity Correction Factor

- Is a method to improve the accuracy of the normal approximation to
- Examine Figure 6.22 from Walpole, Myers, Myers & Ye. Note that each rectangle is centered at x and extends from x-0.5 to x+0.5
- This table should help formulate problems

Binomial Probability	with Correction Factor	Normal Approximation
$P(X \ge x)$	$P(X \ge x - 0.5)$	$P\left(Z > \frac{x - 0.5 - np}{\sqrt{np(1-p)}}\right)$
$P(X \leq x)$	$P(X \le x + 0.5)$	$P\left(Z < \frac{x+0.5-np}{\sqrt{np(1-p)}}\right)$
P(X = x)	$P(x-0.5 \le X \le x+0.5)$	$P\left(Z < \frac{x+0.5-np}{\sqrt{np(1-p)}}\right)$ $P\left(\frac{x-0.5-np}{\sqrt{np(1-p)}} < Z < \frac{x+0.5-np}{\sqrt{np(1-p)}}\right)$







Chapter 4 Content

Normal Approximation to the Poisson Distribution

$$Z = \frac{X - \lambda}{\sqrt{\lambda}}$$

• If X is a Poisson random variable with $E(X) = \lambda$ and $V(X) = \lambda$ $Z = \frac{X - \lambda}{\sqrt{\lambda}}$ is approximately a standard normal random variable.

27 / 42

Exponential Distribution

- The exponential distribution is widely used in the area of reliability and life-test data.
- Ostle, et. al. (1996) list the following applications of the exponential distribution
 - the number of feet between two consecutive erroneous records on a computer tape,
 - the lifetime of a component of a particular device,
 - the length of a life of a radioactive material and
 - the time to the next customer service call at a service desk

Exponential Distribution

 \bullet The PDF for an exponential distribution with parameter $\lambda>0$ is

$$f(x) = \lambda e^{-\lambda x}$$
, for $0 \le x < \infty$

ullet The mean of X is

$$\mu = E(X) = \frac{1}{\lambda}$$

$$\sigma^2 = V(X) = \frac{1}{\lambda^2}$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} dx dx$$

ullet The variance of X is

$$\sigma^2 = V(X) = \frac{1}{\lambda^2}$$

Note that other authors define $f(x) = \frac{1}{\theta}e^{-x/\theta}$. Either definition is acceptable. However one must be aware of which definition is being used.

29 / 42

The Exponential CDF

The CDF for the exponential distribution is easy to derive

$$F(x) = P(X \le x) = \int_{-\infty}^{x} \lambda e^{-\lambda y} dy$$

$$= \int_{0}^{x} \lambda e^{-\lambda y} dy$$

$$= \left(-e^{-\lambda y}\right)\Big|_{y=0}^{x}$$

$$= -e^{-\lambda x} - (-e^{0})$$

$$= -e^{-\lambda x} + 1$$

$$= 1 - e^{-\lambda x}$$

Morelia

Problem 4-79

4-79. The time to failure (in hours) of fans in a personal computer can be modeled by an exponential distribution with $\lambda = 0.0003$.

- (a) What proportion of the fans will last at least 10,000 hours?
- (b) What proportion of the fans will last at most 7000 hours?

Figure 7: image

D(x) /0,000) = 1 - E(10,000) $= 1 - \left[1 - e^{-x(1000)}\right]$ $= e^{-x(.0003)(10000)}$ = .04979

31 / 42

Lack of Memory Property

• The mathematical definition is

$$P(X < t_1 + t_2 | X > t_1) = P(X < t_2)$$

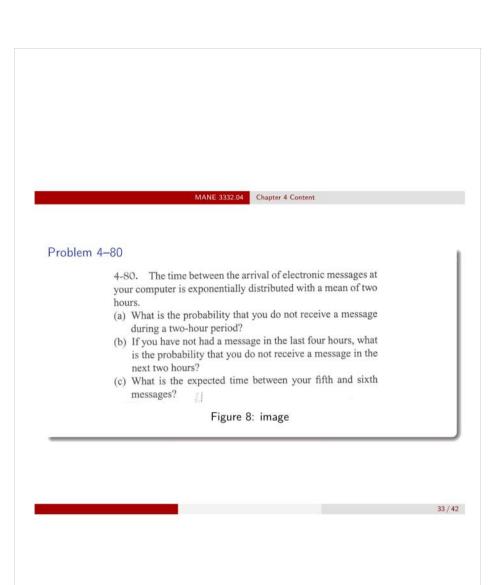
- ullet That is "the probability of a failure time that is less than t_1+t_2 given the failure time is greater than t_1 is the probability that the item's failure time is less than t_2
- This property is unique to the exponential distribution
- Often used to model the reliability of electronic components.

06 Armoche 267,000

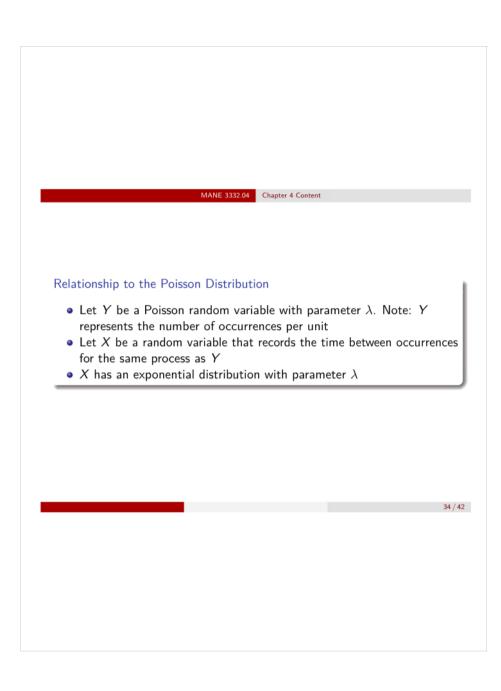
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£2 = 100

New Vehicle L1=100 L2=100



Chapter 4 Page 31



Lognormal Distribution



 \bullet Let W have a normal distribution with mean θ and variance $\omega^2;$ then $X = \exp(W)$ is a **lognormal random variable** with pdf

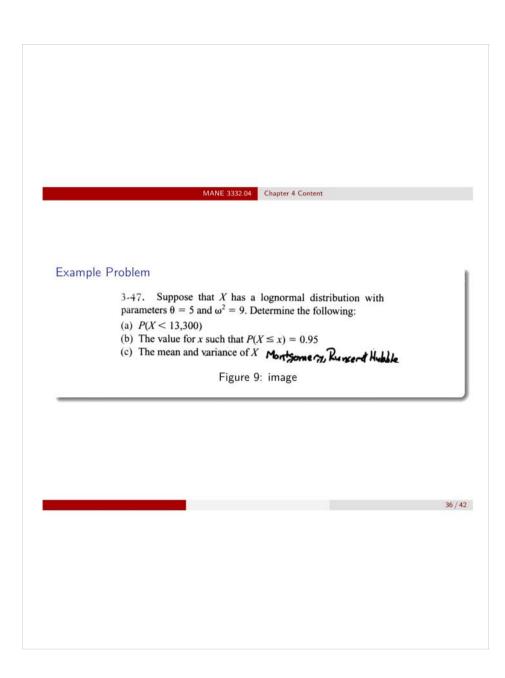
$$f(x) = rac{1}{x\omega\sqrt{2\pi}} \exp\left[-rac{(\ln(x) - heta)^2}{2\omega^2}
ight] \ \ 0 < x < \infty$$

ullet The mean of X is

$$E(X) = e^{\theta + \omega^2/2}$$

 \bullet The variance of X is

$$V(X)=e^{2 heta+\omega^2}\left(e^{\omega^2}-1
ight)$$



- gamma (caps) Chapter 4 Content

Gamma Distribution

• The random variable X with pdf

$$f(x) = \frac{\lambda^r x^{r-1} e^{-\lambda x}}{\Gamma(r)}, \text{ for } x > 0$$

is a **gamma random variable** with parameters $\lambda > 0$ and r > 0.

• The gamma function is

$$\Gamma(r) = \int_0^\infty x^{r-1} e^{-x} dx \text{ for } r > 0$$

with special properties:

• $\Gamma(r)$ is finite • $\Gamma(r) = (r-1)\Gamma(r-1)$ for positive integer r $\Gamma(r) = (r-1)\Gamma(r-1)$ • For any positive integer r $\Gamma(r) = (r-1)\Gamma(r-1)$ • $\Gamma(r) = (r-1)\Gamma(r-1)$

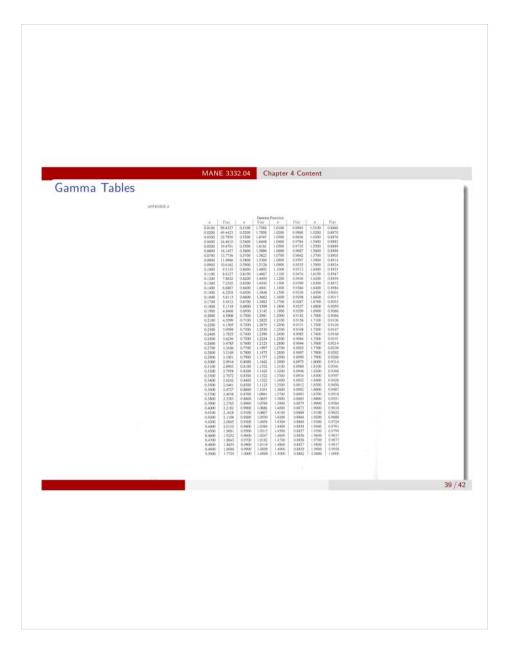
Gamma Distribution

• The mean and variance are

$$\mu = E(X) = r/\lambda$$
 and $\sigma^2 = V(X) = r/\lambda^2$

• We will not work any probability problems using the gamma distribution

38 / 42



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Weibull Distribution

• The random variable X with pdf



$$f(x) = \frac{\beta}{\delta} \left(\frac{x}{\delta}\right)^{\beta-1} \exp\left[-\left(\frac{x}{\delta}\right)^{\beta}\right], \text{ for } x > 0$$

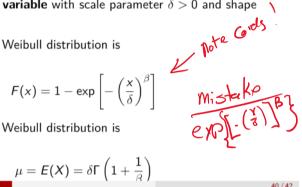
is a **Weibull random variable** with scale parameter $\delta > 0$ and shape parameter $\beta > 0$

• The CDF for the Weibull distribution is



$$F(x) = 1 - \exp\left[-\left(\frac{x}{\delta}\right)^{\beta}\right]$$

• The mean of the Weibull distribution is



$$\mu = E(X) = \delta\Gamma\left(1 + \frac{1}{\beta}\right)$$

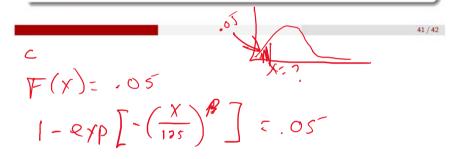
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Weibull Problem

- 45. Suppose that fracture strength (MPa) of silicon nitride braze joints under certain conditions has a Weibull distribution with $\beta = 5$ and $\delta = 125$ (suggested by data in the article "Heat-Resistant Active Brazing of Silicon Nitride: Mechanical Evaluation of Braze Joints," (Welding J., August 1997: 300s-304s).
 - a. What proportion of such joints have a fracture strength of at most 100? Between 100 and 150?
 - b. What strength value separates the weakest 50% of all joints from the strongest 50%?
 - c. What strength value characterizes the weakest 5% of all joints?

Figure 11: image



$$+eyp \left[-\left(\frac{x}{105} \right)^{5} \right] = +.95$$
 $\ln 2exp \left[-\left(\frac{x}{105} \right)^{5} \right] = /n.95$
 $-\left(\frac{x}{105} \right)^{5} = /n.95$

= 1-exp[-(=)

exb[x] = ex

what if t used P X= S(-In (IP)) /B

CDF of uniform

flx)= his for acxeb

Wednesday, February 26, 2025 8:59 AM

$$F(x) = \int_{-\infty}^{\infty} f(x) dx =$$

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$$F(x) = \int f(x) dx = \int \frac{1}{b} dx \qquad for \ x < b$$

$$for \ x > b$$

$$\int_{a}^{b} \frac{1}{b^{2}} dy = \int_{a}^{b} \frac{1}{b^{2}} \frac{1$$

$$c_1 = 49.75$$
 $p(x) > 50 = 1 - F(50)$
 $b = 40.25$ $= 1 - \frac{50 - 49.75}{50.25 - 49.75} = .5$

Monday, March 3, 2025 8:31 AM

Find V Such that P(Z(x)=0.55 5.t. Find value in tibles that has area closest to 0.5 and

receivable of = -7x=0.B

Part g

Monday, March 3, 2025

8:35 AM

1-72=28 .72

Firel x s.t. P(27/x) = .72x = -.58

Monday, March 3, 2025 8:39 AM

QUESTION 1

Let Z be a standard normal random variable, find P(Z>2.65).

- 0.988088
- O The correct answer is not provided.
- 0.011912
- 0.137439
- 0.995975
- 0.975374

0.004025



D(z>265) = 1 - 4265 = 1 - 495975 = .004025

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1-.7825 - .2175 2=? 0

QUESTION 3

Let Z be a standard normal random variable, find the value z such that P(Z>z)=0.7825.

- 0.78
- O The correct answer is not provided.
- 3.12
- O 1.07

0.78

O -1.07

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7-7.0

7 = -.85

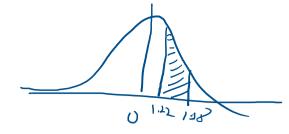
QUESTION 5

Let Z be a standard normal random variable, find the value z such that P(Z < z) = 0.1986.

- 0 1.29
- O The correct answer is not provided.
- 0.85
- 3.1
- O -1.29

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Let Z be a standard normal random variable, find P(1.22<Z<1.28).

- 0.010959
- ____ -0.013696

The correct answer is not provided.

- 0.100273
- 0.888768
- 0.989041

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P(1.22< Z<1.28)= (1.28)- 5(1.22)

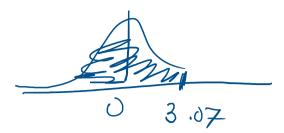
- .899727 - .88876 - .0 | 096

Monday, March 3, 2025 8:56 AM

QUESTION 9

Let Z be a standard normal random variable, find P(Z<3.07).

- 0.003584
- 0.996416
- 0.00107
- O The correct answer is not provided.
- 0.99893
- 0.878258
- 0.839243



D(2<3 D7) = 5(3.07) = .998930

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5.1.3 part b

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First P(X > 11.98)



$$P(x>11.98) = 1 - \frac{11.98 - 10}{2}$$

$$= 1 - \frac{10}{2} (.99)$$

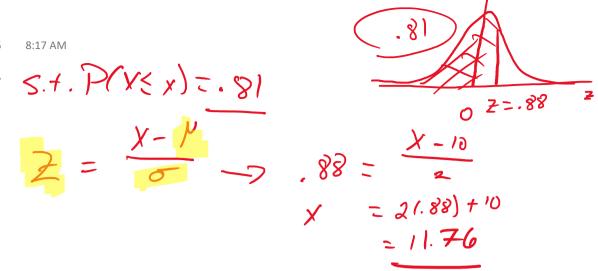
$$= 1 - 838913$$

$$= 1 - 141087$$

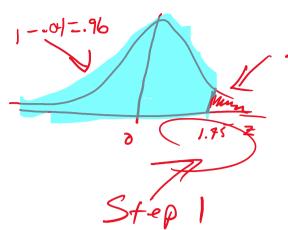
5.1.3 part f

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Fird x st. P(x7,x)= .04



$$\frac{5+ep^2}{2=\frac{x-N}{a}} \rightarrow 1.75 = \frac{x-10}{2}$$

$$1.75 = \frac{X - 10}{2}$$

$$X = 13.5$$

Normal pp

Wednesday, March 5, 2025

QUESTION 1

Let X be a normal random variable with mean (mu) = 109.97 and standard deviation (sigma)=1.054, find P(X<110,195).

- 0.308538
- 0.5
- O The correct answer is not provided.

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- 0.352065
- 0.989879

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$$P(x < 1/0.495) = D(\frac{1/0.495 - 1/09.97}{1.0054})$$

$$= D(0.4978)$$

$$= 0.50$$

$$= 0.50$$

Let X be a normal random variable with mean (mu) = 72.616 and standard deviation (sigma) = 2.581, find the value x such that P(X<x)= 0.8608.

- O 75.403
- O -1.08
 - 0 82.94
 - O The correct answer is not provided.
 - O 69.829



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$$= 1.08 = \frac{X - 72.616}{2.581}$$

X-N

Let X be a normal random variable with mean (mu) = 119.244 and standard deviation (sigma)=1.872, find P(X>117.197).

- The correct answer is not provided.
- 0.841345
 - 0.066844

0.220251

0.862143

0.137857

117.197 119.244 X

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$$P(x7/17.197) = 1 - \frac{1}{5} \left(\frac{1/7.197 - 1/9.244}{1.872} \right)$$

$$= 1 - \frac{1}{3}(-1.09)$$

$$= 1 - .137.857$$

$$= .862.143$$

Let X be a normal random variable with mean (mu) = 104.428 and standard deviation (sigma) = 3.001, find the value x such that P(X>x) = 0.4745.



0 104.248

O The correct answer is not provided.

0 104.608

 $7 = \frac{x - N}{3.001}$

Let X be a normal random variable with mean (mu) = 52.915 and standard deviation (sigma) = 1.553, find P(49.437 < 52.906).

- 0.511966
- 0.499421
- O The correct answer is not provided.
- 0.500579
- 0.366303
- 0.012545

P(49.437<X< 52.966)

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$$= \oint \left(\frac{52.966 - 52.915}{1.553}\right) - \oint \left(\frac{49.437 - 52.5}{1.553}\right)$$

3/5

Attendance 1-1

Grok: Integration by parts

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Sudv= vv-Svde

To solve this, use integration by parts (u=x, $dv=\lambda e^{-\lambda x}dx$):

•
$$du = dx$$

•
$$v = \int \lambda e^{-\lambda x} dx = -e^{-\lambda x}$$

Then:

$$E[X] = \left[x(-e^{-\lambda x})\right]_0^{\infty} - \int_0^{\infty} (-e^{-\lambda x}) dx$$

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ey to intended to 7 e x dy = du

Differelisation 7 x x = u

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QUESTION 1

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Monday, March 10, 2025

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QUESTION 3

Let X be an exponential random variable with lambda=14.134, find the value x such that P(X > x) = 0.846.

- \bigcirc 0.0
- 0.0118
- O 1.0
- 0.1992
- 0.1324

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8:42 AM

QUESTION 5

Let X be an exponential random variable with lambda=11.934, find $P(X \le 8.0E-4)$.

- 0.9905
- O 11.8206
- O -10.8206

0.0095

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Chapter 4 Page 61

Monday, March 10, 2025 8:46 AM

QUESTION 7

Let X be an exponential random variable with lambda=10.367. Find the value x such that P(X < Q)= 0.63.

0.0446

0.0959

0.006

0.9985

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lue x such that
$$P(x < Q) = 0.63$$
.

$$F(x) = .63$$

$$1 - e^{-10.367x} = .63$$

$$+ e^{-10.367x} = +.37$$

$$e^{-10.367x} = -1 \cdot 37$$

Let X be an exponential random variable with lambda=41.432, find P(0.0169 <X< 0.0511).

- O -15.5832
- 0.3761
- 0.4965
- 0.8796

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 $P(.0169 \times (.0511) = F(.0511) - F(.0469)$ $P(.0169 \times (.0511) = F(.0511) - F(.0469)$ $P(x) = 1 - e^{-41.32x} (given)$ $7 = 1 - e^{-41.32(.0169)}$

Let X be a Weibull random variable with parameters delta=5663.526 and beta=2.437. Find the value x such that P(X < x) = 0.065.

- 0 8556.2183
- O 1.0

F(x) = .065 $1 - exp[-(\frac{x}{5663.526})^{0.437}] = .065$

 $(\frac{1}{5}6354)^{2.1137} = (-1n.935)$ $\frac{1}{5}66354)^{2.1137} = (-1n.935)^{2.1137}$ $\frac{1}{5}663.526 = (-1n.935)^{2.1137}$ 1 = 5663.526 (-1n.935) 1 = 1870.3878

Let X be a Weibull random variable with parameters delta=9320.788 and beta=1.284, find the value x such that P(X > x)=0.801.

0.0

1 - [1 - exp[-(x)] - [x] - [

Chapter 4 Page 65

Weibull

Wednesday, March 12, 2025

8:41 AM

QUESTION 7

Let X be a Weibull random variable with delta=10368.214 and beta=4.443, find P(8805.0702 < X < 9320.1677).

F(9320.1677)

- 0.4636
- 0.0
- 0.08
- 0.6164

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exponential CDE $F(X) = 1 - exp[-\lambda]$

exponential is case

exponential is case

| Case | Cure but | S = 1

Weibill CDF

E(X)=1-6xb[-(4)]

F(x) = 1- exp[-(3)]

What if $S = \frac{1}{\lambda}$ $F(x) = 1 - exp \sum_{i=1}^{\infty} \frac{1}{x^i}$