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Section 1

MANE 3332.04

andouts

- Chapter 9 slides
- Chapter 9 slides marked

ıtroduction to Hypothesis Testing

ecision Making for a Single Sample

- Inferential statistics consists of methods used to make decisions or draw conclusions about a population using information contained in a sample
- Inference is divided into two major areas:
  - Parameter estimation (both point and interval)
  - Hypothesis testing

verview of Statistical Hypotheses

- Many engineering problems require a decision to be made regarding some statement about a parameter
  - The statement is called a **hypothesis**
  - The decision-making process about the hypothesis is call **hypothesis testing**
- Statistical hypothesis testing is usually the data analysis stage of a **comparative experiment**
- A procedure leading to a decision about a particular hypothesis is called a **test of hypothesis**
- Testing the hypothesis involves taking a random sample, computing a **test statistic** from the sample data and then using the to make a decision

Null hypothesis

$H_0: \mu = \text{value}$

$H_1: \begin{cases} \mu \neq \text{value} \\ \mu < \text{value} \\ \mu > \text{value} \end{cases}$

Court

H_0: accused is innocent

H_1: Accused is Not innocent

tatistical Hypothesis

- A **statistical hypothesis** is a statement about the parameters of one or more populations
- A statistical hypothesis has two parts a null hypothesis (denoted H_0) and an alternative hypothesis (denoted H_1)
  - The null hypothesis contains an equality statement about the value of parameter. For example $H_0 : \mu = 12$ ounces.
  - There are three possible alternative hypotheses: $H_1 : \mu \neq 12$, $H_1 : \mu < 12$, or $H_1 : \mu > 12$
  - The goal of the research will determine the appropriate alternative hypothesis

Summary of One-Sample Hypothesis-Testing Procedures

Case	Null Hypothesis	Test Statistic	Alternative Hypothesis
1.	$H_0: \mu = \mu_0$ σ^2 known	$z_0 = \dfrac{\bar{x} - \mu_0}{\sigma/\sqrt{n}}$	$H_1: \mu \neq \mu_0$ $H_1: \mu > \mu_0$ $H_1: \mu < \mu_0$

rrors in hypothesis testing

- Whether a correct decision is made depends upon the true nature of H_0 and the decision arrived at.
- A **type I error** occurs when the null hypothesis is true and the outcome of the test is to reject H_0. The probability of a type I error is denoted as α
- A **type II error** occurs when the null hypothesis is false and the outcome of the test is to fail to reject H_0. The probability of a type II error is denoted as β.
- The **power** of a statistical test is the probability rejecting the null hypothesis H_0 when the alternative hypothesis is true. Power $= 1 - \beta$

Figure 9-1 Decision criteria for testing H_0: $\mu =$ 50 centimeters per second versus H_1: $\mu \neq$ 50 centimeters per second.

Table 9-1 Decisions in Hypothesis Testing

Decision	H_0 Is True	H_0 Is False
Fail to reject H_0	no error	type II error
Reject H_0	type I error	no error

rror Example: Manufacturing

rror Example: Medical

eneral Procedure for Hypothesis Testing

he following sequence of steps is recommended

1. From the problem context, identify the parameter of interest,
2. State the null hypothesis, H_0,
3. Specify an appropriate alternative hypothesis, H_1,
4. Choose a significance level α
5. State an appropriate test statistic,
6. State the rejection region for the (test) statistic,
7. Compute any necessary sample quantities, substitute these into the equation for the test statistics, and compute that value,
8. Decide whether or not H_0 should be rejected and report in the problem context

hapter 9, Case 1

ıference on the Mean of a population, variance known

- Assumptions:
  1. $X_1, X_2, \ldots, X_n$ is a random sample of size n from a population
  2. The population is normal, or if it is not normal, the conditions of the central limit theorem apply
- The parameter of interest is μ
- The null hypothesis is $H_0 : \mu = \mu_0$
- The test statistic is

$$Z_0 = \frac{\overline{X} - \mu_0}{\sigma/\sqrt{n}}$$

ıd has a standard normal distribution

- The alternative hypotheses and corresponding critical value(s) are shown in figure 9-11 on page 209

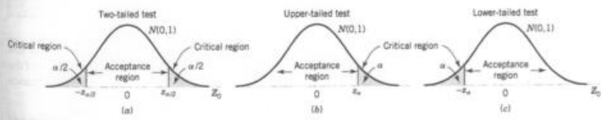

FIGURE 9.11

The distribution of Z_0 when H_0: $\mu = \mu_0$ is true with critical region for (a) the two-sided alternative H_1: $\mu \neq \mu_0$, (b) the one-sided alternative H_1: $\mu > \mu_0$, and (c) the one-sided alternative H_1: $\mu < \mu_0$.

ummary for hypothesis test on the mean, variance known

- See the material on the inside cover of your textbook

nary of One-Sample Hypothesis-Testing Procedures

Null Hypothesis	Test Statistic	Alternative Hypothesis	Fixed Significance Level Criteria for Rejection	P-Value	O.C. Curve Parameter	O.C. Curve Appendix Chart VII
$H_0: \mu = \mu_0$ σ^2 known	$z_0 = \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}}$	$H_1: \mu \neq \mu_0$	$\lvert z_0 \rvert > z_{\alpha/2}$	$P = 2[1 - \Phi(\lvert z_0 \rvert)]$	$d = \lvert \mu - \mu_0 \rvert / \sigma$	a, b
		$H_1: \mu > \mu_0$	$z_0 > z_\alpha$	Probability above z_0 $P = 1 - \Phi(z_0)$	$d = (\mu - \mu_0)/\sigma$	c, d
		$H_1: \mu < \mu_0$	$z_0 < -z_\alpha$	Probability below z_0 $P = \Phi(z_0)$	$d = (\mu_0 - \mu)/\sigma$	c, d

roblem 1

Example 11-1

The burning rate of a rocket propellant is being studied. Specifications require that the mean burning rate must be 40 cm/s. Furthermore, suppose that we know that the standard deviation of the burning rate is approximately 2 cm/s. The experimenter decides to specify a type I error probability $\alpha = 0.05$, and he will base the test on a random sample of size $n = 25$. The hypotheses we wish to test are

$$H_0: \mu = 40 \text{ cm/s},$$
$$H_1: \mu \neq 40 \text{ cm/s}.$$

Twenty-five specimens are tested, and the sample mean burning rate obtained is $\bar{x} = 41.25$ cm/s.

Source: Hines, Montgomery, Goldsman, Borror (2003). Probability and Statistics in Engineering, 4th ed.

roblem 2

9.2.10 The bacterial strain *Acinetobacter* has been tested for its adhesion properties. A sample of five measurements gave readings of 2.69, 5.76, 2.67, 1.62 and 4.12 dyne-cm^2. Assume that the standard deviation is known to be 0.66 dyne-cm^2 and that the scientists are interested in high adhesion (at least 2.5 dyne-cm^2).

a. Should the alternative hypothesis be one-sided or two-sided?

b. Test the hypothesis that the mean adhesion is 2.5 dyne-cm^2.

c. What is the P-value of the test statistic?

ummary Statistics

```
<-c(2.69,5.76,2.67,1.26,4.12)
ibrary(psych)
escribe(x)

#    vars n mean   sd median trimmed  mad  min  max range skew kurtosis   se
# X1    1 5  3.3 1.71   2.69     3.3 2.12 1.26 5.76   4.5 0.26    -1.71 0.76
```

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onnection between Hypothesis Tests and CI

- There is a close connection between confidence intervals and hypothesis tests
- Consider a $100(1-\alpha)\%$ confidence interval on μ and a hypothesis test of size α shown below

$$H_0 : \mu = \mu_0$$
$$H_1 : \mu \neq \mu_0$$

- The conclusion to reject H_0 will be reached if μ_0 is not contained within the confidence interval
- If μ_0 is within the confidence interval, we fail to reject H_0
- The $100(1-\alpha)\%$ confidence interval on μ is the acceptance region

-values

- Is a widely used alternative to the traditional hypothesis test
- Definition: The p-value is the smallest level of significance that would lead to reject of the null hypothesis H_0 with the given data
- Formulas are given below

$$P = \begin{cases} 2[1-\Phi(|z_0|)] & \text{for a two-tailed test} \\ 1-\Phi(z_0) & \text{for a upper-tailed test} \\ \Phi(z_0) & \text{for a lower-tailed test} \end{cases}$$

- Usage: if p-value$<\alpha$ then the conclusion is reject H0, otherwise fail to reject H0 ractice Problem Chapter 9, Case 1 2-sided ractice Problem Chapter 9, Case 1 Lower

ractice Problem Chapter 9, Case 1 Upper

ype II error and sample size for a two-tailed test

- Probability of type II error for the two-tailed test

$$\beta = \Phi\left(z_{\alpha/2} - \frac{\delta\sqrt{n}}{\sigma}\right) - \Phi\left(-z_{\alpha/2} - \frac{\delta\sqrt{n}}{\sigma}\right)$$

here $\mu = \mu_0 + \delta$

- The sample to detect a difference between the true and hypothesized mean of δ with power at least $1 - \beta$ is

$$n \approx \frac{(z_{\alpha/2} + z_\beta)^2\sigma^2}{\delta^2}$$

here $\delta = \mu - \mu_0$

ype II error and sample size for the one-tailed tests

- For an upper-tailed test

$$\beta = \Phi\left(z_\alpha - \frac{\delta\sqrt{n}}{\sigma}\right)$$

- For a lower-tailed test

$$\beta = 1 - \Phi\left(-z_\alpha - \frac{\delta\sqrt{n}}{\sigma}\right)$$

- The sample size required to detect a difference between the true mean and hypothesized mean of δ with power at least $1-\beta$ is

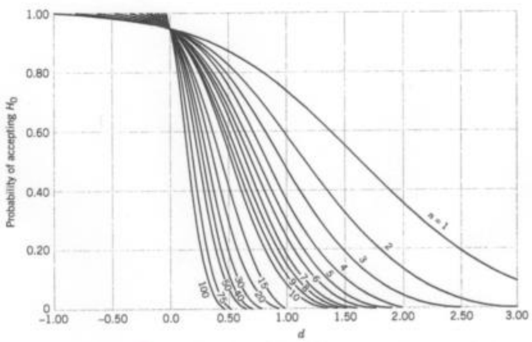
$$n = \frac{(z_\alpha + z_\beta)^2\sigma^2}{\delta^2}$$

n is not an integer, round up to the nearest integer

Z-test

```
library(BSDA)

## Loading required package: lattice

##
## Attaching package: 'BSDA'

## The following object is masked from 'package:datasets':
##
##     Orange

z.test(x,alternative='greater',mu=2.5,sigma.x=0.66,conf.level=0.95)

##
##  One-sample z-Test
##
## data:  x
## z = 2.7104, p-value = 0.00336
## alternative hypothesis: true mean is greater than 2.5
## 95 percent confidence interval:
##  2.814503       NA
## sample estimates:
## mean of x
##       3.3
```

ind the power when the true mean value is 3.325



Power

```
library(asbio)

## Warning: package 'asbio' was built under R version 4.2.3

## Loading required package: tcltk

##
## Attaching package: 'asbio'

## The following object is masked from 'package:psych':
##
##     skew

power.z.test(sigma=0.66,n=5,alpha=0.05,effect=0.825,test="one.tail")

## $sigma
## [1] 0.66
##
## $n
## [1] 5
##
## $power
## [1] 0.8749757
##
## $alpha
## [1] 0.05
##
## $effect
## [1] 0.825
##
## $test
```

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ype II Error Rate and Sample Size

- You will not be required to calculate or use OC-curves
- You must understand the concept and be able to correctly identity type I and type II error

ypothesis Test on the Mean, Variance Unknown

- Much more common case than variance known
- Substitute S for σ
- The test statistics is now a t random variable

$$T = \frac{\overline{X} - \mu_0}{S/\sqrt{n}}$$

ummary of Case 2

$H_0: \mu = \mu_0$ σ^2 unknown	$t_0 = \frac{\bar{x} - \mu_0}{s/\sqrt{n}}$	$H_1: \mu \neq \mu_0$	$\lvert t_0 \rvert > t_{\alpha/2,n-1}$	Sum of the probability above $\lvert t_0 \rvert$ and below $-\lvert t_0 \rvert$	$d = \lvert \mu - \mu_0 \rvert / \sigma$	e, f
		$H_1: \mu > \mu_0$	$t_0 > t_{\alpha,n-1}$	Probability above t_0	$d = (\mu - \mu_0)/\sigma$	g, h
		$H_1: \mu < \mu_0$	$t_0 < -t_{\alpha,n-1}$	Probability below t_0	$d = (\mu_0 - \mu)/\sigma$	g, h

roblem 9.3.6

9.3.6 An article in the *ASCE Journal of Energy Engineering* (1999, Vol. 125, pp. 59–75) describes a study of the thermal inertia properties of autoclaved aerated concrete used as a building material. Five samples of the material were tested in a structure, and the average interior temperatures (°C) reported were as follows: 23.01, 22.22, 22.04, 22.62, and 22.59.

a. Test the hypotheses H_0: $\mu = 22.5$ versus H_1: $\mu \neq 22.5$, using $\alpha = 0.05$. Find the P-value.

b. Check the assumption that interior temperature is normally distributed.

c. Compute the power of the test if the true mean interior temperature is as high as 22.75.

d. What sample size would be required to detect a true mean interior temperature as high as 22.75 if you wanted the power of the test to be at least 0.9?

e. Explain how the question in part (a) could be answered by constructing a two-sided confidence interval on the mean interior temperature.

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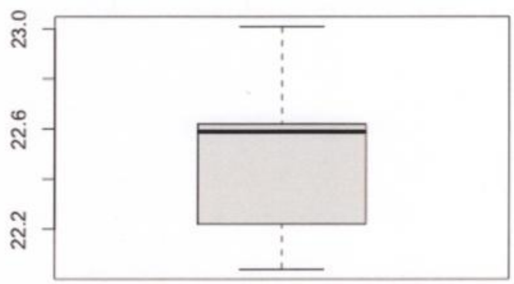
Descriptive Statistics

```
x<-c(23.01,22.22,22.04,22.62,22.59)
library(psych)
describe(x)

##    vars n mean   sd median trimmed  mad   min   max range skew kurtosis
se
## X1    1 5 22.5 0.38  22.59    22.5 0.55 22.04 23.01  0.97 0.08    -1.84
0.17
```

Boxplot

```
boxplot(x)
```



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lassical Approach

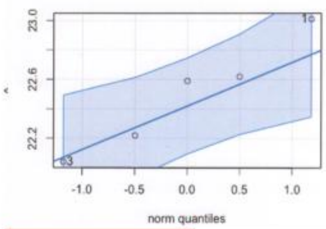
ypothesis Test Using R

t-test

```
t.test(x,alternative="two.sided",mu=22.5,conf.level=0.95)

##
##  One Sample t-test
##
## data:  x
## t = -0.023642, df = 4, p-value = 0.9823
## alternative hypothesis: true mean is not equal to 22.5
## 95 percent confidence interval:
##  22.02625 22.96575
## sample estimates:
## mean of x
##    22.496
```

ormal Probability Plot

```
.brary(car)

# Loading required package: carData
#
# Attaching package: 'car'
# The following object is masked from 'package:psych':
#
#     logit

qPlot(x)
```



'-values

- More difficult to calculate since the t-tables only contain a few quantiles
- Can use tables to generate bounds on the p-value
- Software will provide p-values

-values from R

t-test

```
t.test(x,alternative="two.sided",mu=22.5,conf.level=0.95)
```

```
##
##  One Sample t-test
##
## data:  x
## t = -0.023642, df = 4, p-value = 0.9823
## alternative hypothesis: true mean is not equal to 22.5
## 95 percent confidence interval:
##  22.02625 22.96575
## sample estimates:
## mean of x
##    22.496
```

ower Calculations

- Are much more complicated
- The true distribution is now a non-central t
- Use tables to solve (Chart VII in appendix) or software

ower Calculation using R

ʾower

```
ɔower.t.test(n=5,delta=0.25,sd=0.38,sig.level=0.05,type="two.sample")

##
##      Two-sample t test power calculation
##
##               n = 5
##           delta = 0.25
##              sd = 0.38
##       sig.level = 0.05
##           power = 0.1491624
##     alternative = two.sided
##
## NOTE: n is number in *each* group
```

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ample Size using R

iample Size

```
ower.t.test(power=0.9,delta=0.25,sd=0.38,sig.level=0.05,type="two.sample")

##
##      Two-sample t test power calculation
##
##               n = 49.53305
##           delta = 0.25
##              sd = 0.38
##       sig.level = 0.05
##           power = 0.9
##     alternative = two.sided
##
## NOTE: n is number in *each* group
```

hapter 9, Case 2 2-sided Practice Problems

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hapter 9, Case 2 Upper Practice Problems

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ase 3. Hypothesis Test on Variance of Normal Population

- The test statistics is a χ^2 random variable

$$\chi_0^2 = \frac{(n-1)S^2}{\sigma_0^2}$$

Tests on the Variance of a Normal Distribution

Null hypothesis: $H_0: \sigma^2 = \sigma_0^2$

Test statistic: $\chi_0^2 = \frac{(n-1)S^2}{\sigma_0^2}$

Alternative Hypothesis	Rejection Criteria
$H_1: \sigma^2 \neq \sigma_0^2$	$\chi_0^2 > \chi^2_{\alpha/2,n-1}$ or $\chi_0^2 < \chi^2_{1-\alpha/2,n-1}$
$H_1: \sigma^2 > \sigma_0^2$	$\chi_0^2 > \chi^2_{\alpha,n-1}$
$H_1: \sigma^2 < \sigma_0^2$	$\chi_0^2 < \chi^2_{1-\alpha,n-1}$

The table below summarizes the three possible hypothesis tests. The jection regions are clearly shown in Figure 9-17 on page 222 igure 9-17

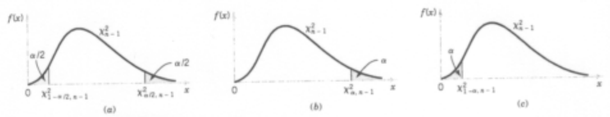


FIGURE 9.17

Reference distribution for the test of H_0: $\sigma^2 = \sigma_0^2$ with critical region values for (a) H_1: $\sigma^2 \neq \sigma_0^2$, (b) H_1: $\sigma^2 > \sigma_0^2$, and (c) H_1: $\sigma^2 < \sigma_0^2$.

est Summary

- See summary in your textbook

$H_0: \sigma^2 = \sigma_0^2$	$\chi_0^2 = \frac{(n-1)s^2}{\sigma_0^2}$	$H_1: \sigma^2 \neq \sigma_0^2$	$\chi_0^2 > \chi^2_{\alpha/2,n-1}$ or $\chi_0^2 < \chi^2_{1-\alpha/2,n-1}$	See text Section 9.4.	$\lambda = \sigma/\sigma_0$	i, j
		$H_1: \sigma^2 > \sigma_0^2$	$\chi_0^2 > \chi^2_{\alpha,n-1}$		$\lambda = \sigma/\sigma_0$	k, l
		$H_1: \sigma^2 < \sigma_0^2$	$\chi_0^2 < \chi^2_{1-\alpha,n-1}$		$\lambda = \sigma/\sigma_0$	m, n

roblem 7.108

roblem taken from Ostle, Turner, Hicks and McElrath (1996). *Engineering tatistics: The Industrial Experience.* Duxbury Press.

.108 Incoming coal at a coking plant is routinely analyzed for sulfur content (in percent). In the past, samples taken from barges loaded with coal from a particular mine have had a variance of 0.000196. When a new analyst was hired, the results of an assay of coal from the mine produced percentages of 0.83, 0.79, 0.77, 0.81, and 0.80.

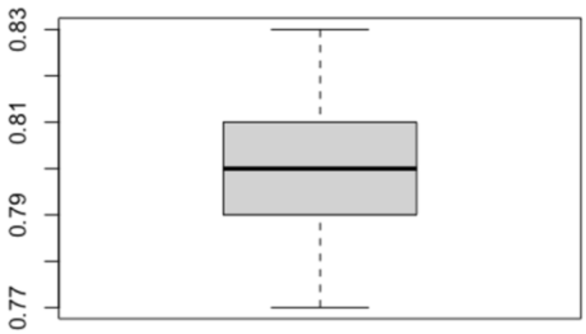
(a) Using $\alpha = 0.05$, does the sample variance provide sufficient evidence to conclude that the results from the new analyst indicate more variability than in the past? State all assumptions.

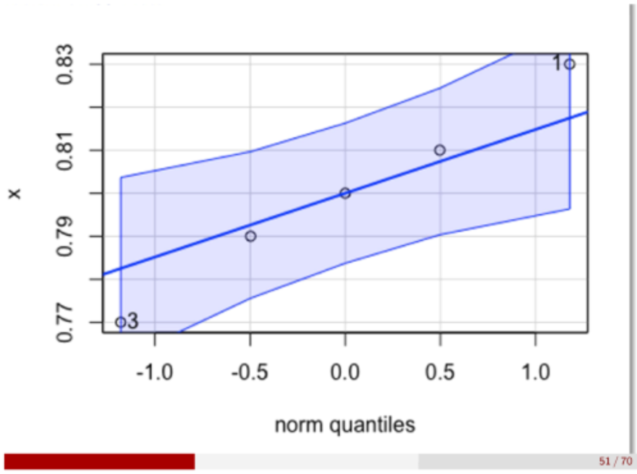
(b) Based on these data, is an assumption of normality reasonable? Justify by using a normal quantile plot *and* a formal test such as the Shapiro-Wilk W test.

tatistics for Problem 7.108

```
<-c(0.83,0.79,0.77,0.81,0.80)
ibrary(psych)
escribe(x)

#    vars n mean   sd median trimmed  mad  min  max range skew kurtosis   se
# X1    1 5  0.8 0.02    0.8     0.8 0.01 0.77 0.83  0.06    0    -1.69 0.01

rint(var(x))

# [1] 5e-04
```

roblem 7.108 Plots




0.83
0.81
0.79
0.77
x
1
3
-1.0
-0.5
0.0
0.5
1.0
norm quantiles
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roblem 7.108 Shapiro-Wilks Test

```
# [1] 3 1

hapiro.test(x)

#
#  Shapiro-Wilk normality test
#
# data:  x
# W = 0.99929, p-value = 0.9998
```

-values

- Very similar to the case for the mean of a normal population with variance unknown
- Difficult to calculate since the χ^2-tables only contain a few quantiles
- Can use tables to generate bounds on the p-value
- Software will provide p-values

```
arTest(x,alternative="greater",conf.level=0.95,sigma.squared=0.000196)

#
# Results of Hypothesis Test
# --------------------------
#
# Null Hypothesis:                 variance = 0.000196
#
# Alternative Hypothesis:          True variance is greater than 0.000196
#
# Test Name:                       Chi-Squared Test on Variance
#
# Estimated Parameter(s):          variance = 5e-04
#
# Data:                            x
#
# Test Statistic:                  Chi-Squared = 10.20408
#
# Test Statistic Parameter:        df = 4
#
# P-value:                         0.03712675
#
# 95% Confidence Interval:         LCL = 0.0002107986
#                                  UCL =          Inf
```

ower Calculations

- Can be done with OC curves found in Table VII*i*–*n*
- Can be done in software such as R

est on Standard Deviation

- What about test on standard deviation?

hapter 9, Case 3 2-sided Practice Problems ractice Problem Chapter 9, Case 3 Lower

ractice Problem Chapter 9, Case 3 Upper

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ase 4. Hypothesis Test on a Population Proportion

- The test statistics for the hypothesis test is

$$Z_0 = \frac{x - np_0}{\sqrt{np_0(1 - p_0)}}$$

$H_0: p = p_0$	$z_0 = \frac{x - np_0}{\sqrt{np_0(1-p_0)}}$	$H_1: p \neq p_0$	$\lvert z_0 \rvert > z_{\alpha/2}$	$p = 2[1 - \Phi(z_0)]$	3–4	3–4
		$H_1: p > p_0$	$z_0 > z_\alpha$	Probability above z_0 $p = 1 - \Phi(z_0)$	3–4	3–4
		$H_1: p < p_0$	$z_0 < -z_\alpha$	Probability below z_0 $P = \Phi(z_0)$	3–4	3–4

roblem 9.5.2

9.5.2 WP Suppose that of 1000 customers surveyed, 850 are satisfied or very satisfied with a corporation's products and services.

a. Test the hypothesis H_0: $p = 0.9$ against H_1: $p \neq 0.9$ at $\alpha = 0.05$. Find the P-value.

b. Explain how the question in part (a) could be answered by constructing a 95% two-sided confidence interval for p.

roblem 9.5.2 Classic Approach

- For the two-sided alternative hypothesis

$$\beta = \Phi\left(\frac{p_0 - p + z_{\alpha/2}\sqrt{p_0(1-p_0)/n}}{\sqrt{p(1-p)/n}}\right) - \Phi\left(\frac{p_0 - p - z_{\alpha/2}\sqrt{p_0(1-p_0)/n}}{\sqrt{p(1-p)/n}}\right)$$

- If the alternative is $H_1 : p < p_0$

$$\beta = 1 - \Phi\left(\frac{p_0 - p - z_{\alpha}\sqrt{p_0(1-p_0)/n}}{\sqrt{p(1-p)/n}}\right)$$

- and finally if the alternative hypothesis is $H_1 : p > p_0$

$$\beta = \Phi\left(\frac{p_0 - p + z_{\alpha}\sqrt{p_0(1-p_0)/n}}{\sqrt{p(1-p)/n}}\right)$$

ample Size

- Sample size requirements to satisfy type II(β) error constraints for a two-tailed hypothesis test is given by

$$n = \left[\frac{z_{\alpha/2}\sqrt{p_0(1-p_0)} + z_\beta\sqrt{p(1-p)}}{p - p_0}\right]^2.$$

- For a sample size for a one-sided test substitute z_α for $z_{\alpha/2}$.
- Problem 9.95

esting for Goodness of Fit

- Material is presented in section 9-7 of your textbook
- Procedure determines if the sample data is from a specified underlying distribution
- Procedure uses a χ^2 distribution
- Example 9-12 presents a χ^2 goodness of fit test for a Poisson example
- Example 9-13 presents a χ^2 goodness of fit test for a normal example

rocedure

1. Collect a random sample of size n from a population with an unknown distribution,
2. Arrange the n observations in a frequency distribution containing k classes
3. Calculate the observed frequency in each class O_i,
4. From the hypothesized distribution, calculate the expected frequency in class i, denoted E_i (if E_i is small combine classes)
5. Calculate the test statistic

$$\chi_0^2 = \frac{\sum_{i=1}^{k} (O_i - E_i)^2}{E_i}$$

6. Reject the null hypothesis if the calculated value of the test statistic $\chi_0^2 > \chi^2_{\alpha, k-p-1}$ where p is the number of parameters in the hypothesized distribution

xample 9.12, part 1

EXAMPLE 9.12 | Printed Circuit Board Defects—Poisson Distribution

The number of defects in printed circuit boards is hypothesized to follow a Poisson distribution. A random sample of $n = 60$ printed circuit boards has been collected, and the following number of defects observed.

Number of Defects	Observed Frequency
0	32
1	15
2	9
3	4

The mean of the assumed Poisson distribution in this example is unknown and must be estimated from the sample data. The estimate of the mean number of defects per board is the sample average, that is, $(32 \cdot 0 + 15 \cdot 1 + 9 \cdot 2 + 4 \cdot 3)/60 = 0.75$. From the Poisson distribution with parameter 0.75, we may compute p_i, the theoretical, hypothesized probability associated with the ith class interval. Because each class interval corresponds to a particular number of defects, we may find the p_i as follows:

$$p_1 = P(X = 0) = \frac{e^{-0.75}(0.75)^0}{0!} = 0.472$$

$$p_2 = P(X = 1) = \frac{e^{-0.75}(0.75)^1}{1!} = 0.354$$

$$p_3 = P(X = 2) = \frac{e^{-0.75}(0.75)^2}{2!} = 0.133$$

$$p_4 = P(X \geq 3) = 1 - (p_1 + p_2 + p_3) = 0.041$$

The expected frequencies are computed by multiplying the sample size $n = 60$ times the probabilities p_i. That is, $E_i = np_i$. The expected frequencies follow:

Number of Defects	Probability	Expected Frequency
0	0.472	28.32
1	0.354	21.24
2	0.133	7.98
3 (or more)	0.041	2.46

Because the expected frequency in the last cell is less than 3, we combine the last two cells:

Number of Defects	Observed Frequency	Expected Frequency
0	32	28.32
1	15	21.24
2 (or more)	13	10.44

The seven-step hypothesis-testing procedure may now be applied, using $\alpha = 0.05$, as follows:

1. **Parameter of interest:** The variable of interest is the form of the distribution of defects in printed circuit boards.
2. **Null hypothesis:** H_0: The form of the distribution of defects is Poisson.
3. **Alternative hypothesis:** H_1: The form of the distribution of defects is not Poisson.
4. **Test statistic:** The test statistic is $\chi_0^2 = \sum_{i=1}^{k} \frac{(O_i - E_i)^2}{E_i}$
5. **Reject H_0 if:** Because the mean of the Poisson distribution was estimated, the preceding chi-square statistic will have $k - p - 1 = 3 - 1 - 1 = 1$ degree of freedom. Consider whether the P-value is less than 0.05.
6. **Computations:**

$$\chi_0^2 = \frac{(32 - 28.32)^2}{28.32} + \frac{(15 - 21.24)^2}{21.24} + \frac{(13 - 10.44)^2}{10.44}$$
$$= 2.94$$

7. **Conclusions:** We find from Appendix Table III that $\chi^2_{0.10,1} = 2.71$ and $\chi^2_{0.05,1} = 3.84$. Because $\chi_0^2 = 2.94$ lies between these values, we conclude that the P-value is between 0.05 and 0.10. Therefore, because the P-value exceeds 0.05, we are unable to reject the null hypothesis that the distribution of defects in printed circuit boards is Poisson. The exact P-value computed from software is 0.0864.

hapter 9 Summary

- You should be prepared to work any practice problems assigned: Cases 1-3 with three different alternatives
- All other information is conceptual knowledge that can be questioned with multiple choice
  - Name 3 ways to test if data is from a normal distribution